Northwestern Northwestern

IPR Working Paper Series

WP-22-13

Social Disruption, Gun Buying, and Anti-System Beliefs

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Version: March 15, 2022

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Abstract

Gun ownership is a highly a consequential political behavior. It often signifies a belief about the inadequacy of state-provided security and leads to membership in a powerful political constituency (which is commonly mobilized by the National Rifle Association). As such, understanding why people purchase guns and how doing so affects the composition of gun owners is important, as it can have palpable political consequences. The researchers address these issues by exploring the dynamics of one of the largest gun-buying spikes in American history, which took place during the COVID-19 pandemic. They show that feelings of diffuse threat prompted many to buy guns. Moreover, new gun owners, even more than buyers who already owned guns, exhibit strong conspiracy and anti-system beliefs. This has substantial consequences for the subsequent population of gun owners, and provides insight into how social disruptions can alter the nature of political groups.

The authors thank Mark Joslyn and Mike Miller for helpful comments. They acknowledge financial support from the National Science Foundation under grants SES-2029292, SES-2029297, and SES-2116645, and the Peter G. Peterson Foundation.

Groups are central to politics, with competition and cooperation among them playing an important role in, among other things, shaping public policy outcomes and structuring partisan politics. Indeed, several foundational models of American democracy place groups front and center, a view expressed as early at the 1830s when de Tocqueville commented on the propensity of Americans to form "associations" (Tocqueville 1835). While these models contest the extent to which group competition in the U.S. system promotes relatively equal democratic representation, they nonetheless all tend to portray groups as the key building blocks of politics.¹ One stream of this conceptualization focuses on the effects of "disturbances," exploring how system-disrupting developments can cause individuals to feel threatened, thereby prompting groups into action. This idea spans generations of work, beginning with Truman's (1951) argument that "potential" groups form due to social, economic, and political disturbances, extending to Hansen's (1985) point that individuals are more apt to join groups during threatening times, and continuing with (among many others) Klar's (2013) finding that messages involving identify.

Here, we utilize and extend this framework to study a remarkable phenomenon set off by recent disturbances: a massive surge in gun purchases that began immediately following the outbreak of COVID-19 and continued through a summer of protests and a tumultuous election cycle. During 2020, a record-breaking 17 million Americans purchased one or more firearms, a spike which first shows up in FBI background check data for April 2020, shortly after COVID-19 had established itself in all 50 states (Denham and Ba Tar 2021; Fisher et al. 2021; Tavernise 2021). This was a year that witnessed one of the most large-scale disturbances of the last

¹ See, among others, Truman (1951) and Dahl (1961) for optimistic takes about pluralist democracy and Schattschneider (1960) and Olson (1965) for more critical takes. Despite their intense debates about the consequences of group-based politics for democratic representation, it is notable that both "pluralists" and their critics generally agreed about the central role of groups in politics.

century—the COVID-19 pandemic—as well historic protest events and notable political turmoil, which together had dramatic health, economic, and social consequences for the country. While these disturbance-driven gun purchases raise numerous questions, our focus here is on their consequences for the future of the gun-owning community. How might the momentous gun-buying surge of 2020 durably transform the composition, politics, and demands of gun owners as a group?

The answer to this question has important consequences. Indeed, buying a gun often signifies an individual's desire to address security concerns on their own rather than relying on the state, which—especially when events in the world make individuals feel unsafe and gun ownership grows—can have major sociopolitical consequences. Moreover, gun owners constitute a crucial and unusually politically-engaged group: many gun owners share a highly politically-salient social identity, which has been central to the mobilizational capabilities of the historically powerful National Rifle Association (NRA) and helped cement its prominent position in right-wing politics (Lacombe 2019, 2021; Joslyn 2020; Lacombe et al. 2019; Joslyn and Haider-Markel 2017). With the NRA experiencing substantial organizational challenges that pre-date the disturbances of 2020, the year's gun-buying surge has the potential to either undermine or buttress it moving forward.

For all of these reasons, there has been speculation about how 2020 may have changed (or not changed) the gun-owning community; much of this commentary has focused on anecdotal reports of increased gun buying among Black Americans and women, demographics not typically associated with gun ownership. If such trends are indeed borne out by the data, one

implication could be that new gun owners will politically moderate the population of gun owners.² Whether this is the case, however, is yet to be seen—and, in fact, we suggest otherwise.

We theorize that the disturbances of 2020 and into 2021³ generated threats that motivated many individuals to purchase guns for the first time. As a result of having been disproportionately motivated by feelings of threat, we argue that the year's new gun buyers are compositionally distinct from pre-existing gun owners; that is, whereas the latter includes gun hobbyists (e.g., hunters and target shooters) *and* individuals motivated by threat, the former is disproportionately comprised of those driven by threat. This matters because threat has *also* been linked to an increased likelihood of holding conspiratorial and anti-system beliefs. Consequently—and contrary to the speculation mentioned earlier—we expect that 2020's new gun buyers are more likely to hold such beliefs than individuals who already owned firearms, and have therefore altered the shape of the gun-owning community to include more people suspicious of the system.

We test our expectations with a large survey of more than 7,000 gun owners, in which we differentiate first-time and pre-existing owners, and examine their views across several relevant outcomes. Our results confirm our expectations: new gun owners are more likely to hold conspiracy beliefs and less likely to trust governmental institutions. We also show that gun buying in general during 2020 correlates significantly with diffuse threat variables such as having COVID-19 in one's household and economic hardship. Overall, our findings contradict extant narratives, discussed above, that the gun buying spike of 2020 might moderate the population of gun buyers. Indeed, we find that new gun owners' views differ from those of pre-

² See, for example, Paterson (2020), Young et al. (2021), Alcorn (2020), NPR (2021), O'Rourke (2020), Yamane (2021), and Linthicum (2020).

³ In the remainder of the paper, references to "2020" include the period from March 2020 through the collection of our data in April 2021.

existing gun owners, but this shift moves the views of the group as a whole in a more, not less, extreme direction. The shift we identify has palpable implications for democracy given that gun owners, as a group, have the means and a now increased motivation to act violently against the state—or against fellow citizens whom they associate with it. More broadly, our results provide an example of how major social disturbances can affect groups not only by altering opinions or mobilizing actions, but also by changing the composition of their membership in politicallyimportant ways.

Understanding Gun Purchases During Threatening Times

Threat is a powerful stimulant of action. When people feel threatened, they often become anxious and respond in ways that they believe can minimize danger (e.g., Reiss et al. 2021). This has been demonstrated in individuals' attitudes across multiple domains, including climate change (Stollberg and Jones 2021), terrorism (Sloan et al. 2021), penal response (Klar 2013), personal health (Horner et al. 2021), and more. Most relevant to our paper, purchasing a gun is also a documented response to threat. Sloan et al. (2021), for example, show that fear of Muslim terrorist attacks increases the likelihood of buying a firearm. More generally, Stroebe et al. (2017) offer a theory of gun purchasing that posits the impact of both specific threats—such as victimization—as well diffuse threats that come from a belief that the world is dangerous and unpredictable (also see Warner and Thrash 2019). These latter types of threats induce fear that causes unease about the social order (Jackson 2006). Along these lines, Warner (2020: 12), in her study of the motivations of gun ownership, states that in general "fear of crime [is] rooted more broadly in abstract anxieties about modernization, reflecting diffuse anxieties brought on by social and economic changes, and perceptions of the world as chaotic and out of control." This

coheres with Carlson's (2015) finding that gun carriers conflate crime and economic decline. These diffuse threats can lead to gun buying in order to gain a sense of protection, even if the purchasers do not consciously identify the source of anxiety, such as whether it concerns crime, economic challenges, or some other source (Warner 2020).

These types of sentiments likely help explain the unprecedented spike in gun purchases that occurred during 2020. In fact, from the perspective of the work cited above, it is somewhat unsurprising that the generally threatening atmosphere experienced by Americans in 2020 led to gun buying. The pandemic introduced a range of novel threats—including health threats from the virus itself, social threats from isolation, and economic threats due to widespread hardship (e.g., Perlis et al. 2021)—which together (and in conjunction with protest events and political turmoil) appear to have motivated individuals to purchase firearms; indeed, given the variety and magnitude of the threats Americans faced, the fact that increased firearm background checks in 2020 dwarfed prior gun-buying events rather clearly reflects a perception of guns as a source of safety from a broad sense of peril (Lang and Lang 2021; also see Kerner et al. 2022). Therefore, to say that the initial gun-buying surge in April 2020 (and among households sick with COVID-19) was fueled by the pandemic is not to imply that purchasers bought a gun to fight a virus. Rather, fear of death and illness—combined with layoffs, lockdowns, and shortages of essential household products—created a generalized anxiety which in turned fueled gun buying.

Extending this line of thinking, we argue that a surge in threat-motivated gun purchases—especially of the size that occurred in 2020—will affect the composition of gun owners as a group in important ways. To see why, consider four points: First, during less troubling times, a non-trivial number of people buy guns for reasons orthogonal to threat, most notably for hunting and target shooting; a 2017 poll, for example, found that 38% of gun owners

reported hunting and 30% sport shooting as their reasons for ownership (Parker et al. 2017).⁴ Second, as explained above, we expect that during COVID-19, feelings of threat likely played an outsized role in gun purchases. Those who bought guns during COVID-19, we argue, often did so due to a diffuse sense of threat brought on by economic, health, and other concerns. Third, if threats induced by 2020 played a large role in motivating gun purchases, it then follows that the group of individuals who bought guns *for the first time* during COVID-19 will, compositionally, be comprised of a greater proportion of individuals who were motivated by threat than the larger population of gun owners (a substantial proportion of whom have bought firearms, at least in part, for hobbyist reasons). Fourth, these new gun buyers—because their purchases were motivated by threat to an unusual degree—will then be more likely than prior gun owners to hold other beliefs (discussed below) that are correlated with threat. As a result, the arrival of these new gun owners into the gun owning community will alter the overall composition of beliefs within that community moving forward.

These "other" beliefs include those related to conspiracies—that is, beliefs that seek to explain an event by invoking the machinations of powerful people, who attempt to conceal their role while pursuing malevolent goals (Bale 2007; Sunstein and Vermeule 2009). Conspiracy ideation comes in many guises; for example, believing that NASA faked the moon landing, or that the government suppressed evidence that the MMR vaccine causes autism. When individuals experience threatening events that induce anxiety (e.g., natural disasters, disease outbreaks), they feel a loss of control (Landau et al. 2015, van Prooijen and Douglas 2017, van Prooijen 2019), which in turn increases their likelihood of accepting conspiracy theories. As

⁴ To be sure, protection constituted an important reason for gun buying prior to the pandemic (Yamane 2017). Our argument (which we provide evidence for in the next section of the paper) is not that protection-motivated gun buying is a new phenomenon, but instead that pandemic gun buyers were motivated by protection to a greater extent than past gun buyers.

Strol et al. (2021: 721) put it, individuals "take a complex event—for example, an outbreak of a deadly virus—and provide an explanation of the event and someone to blame for it…" which indicates that "conspiracy theories may satisfy important epistemic motives, that is, the need to understand what is happening around us, as well as existential motives to regain the feeling of control, security, and meaning in the world after encountering some threatening event." Related to 2020's events, Strol et al. (2021) show that perceptions of COVID-19 risk and a concomitant lack of control predicts COVID-19 conspiracy beliefs, as well as more generic conspiracy and pseudoscientific beliefs (i.e., it is not domain specific) (also see Jutzi et al. 2020, Scrima et al. 2022).

We use these findings to posit a distinction between new and old gun owners. Since the group of new gun buyers will be composed of more individuals driven by threat, they will also be more likely to accept both general and COVID-19 specific conspiracy theories. Our first hypothesis, then, is as follows.

Gun buyers who purchased firearms for the first time during 2020-21 will be significantly more likely than those who previously owned guns to hold general and specific COVID-19 conspiracy beliefs, all else constant. (Hypothesis 1)

Threat also relates to anti-system beliefs and trust. When citizens attribute a threatening situation to governmental actors, their trust in those actors declines—they are unable to trust those who they see as having caused the threat (e.g., Albertson and Gadarian 2015, Schlipphak 2021). This explains why partisans who particularly dislike or feel threatened by the other party become distrustful of government when that party wins office (Hetherington and Rudolph 2015).

Building on our prior point, if more first-time gun buyers bought due to threat, it follows that they will express less trust in institutions than those who already owned them. Similarly, the increased conspiracy beliefs among these new gun owners (as suggested by hypothesis 1) means they likely have less faith in institutions (i.e., they attribute institutional failure as a source of the threat). In our case, this includes health and scientific institutions (which may be seen as having failed to adequately handle COVID-19) as well as media institutions (which may be seen as having mislead the public about the pandemic and other relevant matters). This leads to our second hypothesis.

Gun buyers who purchased firearms for the first time during 2020-21 will be significantly less trusting of health, scientific, and media institutions than those who previously owned guns, all else constant. (Hypothesis 2)

Importantly, our hypotheses, if confirmed, would be all the more notable given that the population of pre-existing gun owners would themselves be expected (relative to the general public) to hold the sorts of views described above. That is, our hypotheses about the attitudes of new gun owners are not meant to imply that pre-existing owners are *unlikely* to hold conspiratorial views about 2020's events and anti-system sentiments about actors and institutions who played central roles in those events. In fact, given their notably strong support for the rightwing populist worldview popularized by Trumpism and long-promoted by the NRA (Lacombe 2021), we expect (all else equal) pre-existing gun owners, when compared to other Americans, to be significantly more likely to hold such views. Our theoretical framework and associated hypotheses, however, lead us to expect that first-time gun owners will shift the broader gun-

owning community *even further* in this direction, reinforcing and extending the sorts of extant attitudes that have been shown to be associated with gun ownership.

Data and Methods

Our data come from a large, online survey, with participants recruited via PureSpectrum between April and July of 2021.⁵ Our full sample includes 24,448 individuals and has been weighted to reflect the U.S. population along dimensions of race/ethnicity, gender, age, education, geographic region, and county urbanicity. (In Appendix A, we provide a table containing descriptive statistics of the sample.) Our analyses that compare gun owners to other Americans use the full sample, while our analyses comparing pre-existing and new gun owners focus on a subset of respondents (N=7,699) who reported being gun owners. We consider pandemic (or "2020") gun buyers to be those who bought guns in or after March 2020—when COVID-19's presence in the U.S. began to rapidly increase, a national emergency was declared, and states throughout the country issued stay-at-home orders. Pre-existing gun owners are those who—regardless of whether they made pandemic purchases—owned guns prior to March 2020, while first-time buyers are those who bought guns during or after March 2020 and did not, prior to that point, own any.

We begin by assessing the relationship between threat and 2020 gun buying. As noted earlier, there are clear reasons to suspect that pandemic gun buyers were motivated by the different types of threats that were caused by the events of 2020. To examine this, we use as our dependent variable a question asking respondents whether they or a member of their household purchased a gun during the pandemic (see Appendix B for exact wording). We use linear

 $^{^{5}}$ Unless otherwise noted, all results are pooled from two waves. The April 2021 wave lasted 4/1-5/3 while the June wave lasted 6/9-7/6. If a respondent participated in both waves, only their first observation was retained.

probability models (with robust standard errors) to examine the effect of a number of factors on gun purchasing.⁶ More specifically, as our primary independent variables of interest, we use two different measures to capture factors that would produce the sort of diffuse sense of threat related to the pandemic and the year's events discussed earlier: first, whether they report having experienced economic hardships during the pandemic and, second, whether anyone in an individual's household was diagnosed with COVID-19. As a point of clarification, we do not mean to suggest that individuals consciously or explicitly connected these experiences to gun purchasing; as explained, rather, they generate a sense of diffuse threat (or anxiety) that leads one to take action in response. (We later will offer some insight into individuals' explicitly offered rationales for buying guns.)

We also include a range of other variables that might affect pandemic gun purchasing, including partisanship, parental status, race, community type (rural, urban, or suburban), whether the respondent is a White evangelical Christian (which has been shown to be linked to gun ownership and attitudes; Yamane 2016; Merino 2018), income, ideology, college education, gender, and age. We also control for prior gun ownership (which, as a predictor of future gun purchases, is important to hold constant in order to identify the impact of threat), and region fixed effects.

In this first set of tests, we expect to find that pandemic gun buying is predicted by each of our variables capturing threat (i.e., economic hardships and household COVID diagnoses); to

⁶ We report linear probability and ordinary least squares models throughout the paper rather than maximumlikelihood models, as they require fewer assumptions and modeling decisions while typically producing substantively similar results and allowing for easier interpretation. Nonetheless, as a check, we also estimated logit models (see Appendix C); our results hold under these alternative specifications. See Angrist and Pischke (2009: 102–7).

be clear, we expect this to be the case for all individuals, including those who did and did not own guns prior to the pandemic.

We then shift to analyses that examine our core hypotheses regarding conspiracy beliefs and institutional trust. In this part of the paper, we compare new and pre-existing gun owners to examine the extent to which their views differ. We do this in two different ways; first, by comparing new gun owners to all pre-existing gun owners (regardless of whether those preexisting owners bought additional guns during the pandemic) and, second, by comparing new gun owners to pre-existing owners who did *not* buy more guns during the pandemic. We also, as a point of reference, compare pre-existing gun owners to non-gun owners (i.e., those who did not own guns before the pandemic and did not buy them during it), which provides a baseline measure of the views of those who owned guns prior to the pandemic. These latter analyses give important context, as they speak to where gun owners as a social group stood prior to the entry of first-time buyers into the gun owning community; the substantive consequences of differences and/or similarities between the views of new and pre-existing gun owners depends on the nature of pre-existing gun owners' views. In other words, the consequences of our main findings which compare new gun owners to pre-existing owners-depend in part on how likely preexisting owners are to hold conspiracy beliefs and how trusting they are of institutions.

To test our first hypothesis, we look at two dependent variables that measure conspiracy beliefs. The first captures whether individuals believe that the 2020 election was stolen from Donald Trump (see Graham and Yair 2021 regarding the depth and stability of this belief as reported on surveys). The second is an additive index capturing conspiratorial views about COVID-19 vaccines; this consists of 5 items pertaining to whether the respondent believes that vaccines change people's DNA, contain microchips, incorporate lung tissue from aborted

fetuses, or cause infertility.⁷ We also included a true item (that the vaccine has been thoroughly tested), reverse coded such that the variable takes on the value of 1 if the respondent does *not* indicate it is true that the COVID-19 vaccines were tested on thousands of people in clinical trials.

To test our second hypothesis, we examine several dependent variables pertaining to trust in institutions. The goal is to identify whether first-time buyers have less faith in the "system" then pre-existing gun owners. More specifically, we examine trust in health institutions (consisting of the FDA, CDC, and Dr. Anthony Fauci of the NIH combined into an index⁸), scientists, and the news media. (See the Appendix B for exact question wordings.) Beyond their obvious relevance to the events of 2020, we believe our variables pertaining to conspiracy beliefs and trust are particularly useful for testing our hypotheses because they are not directly related to gun politics; in other words, rather than looking at trust in, for example, the Bureau of Alcohol, Tobacco, and Firearms (which enforces most federal gun laws) or beliefs in conspiracies pertaining to gun confiscation, we instead examine outcomes that constitute a tougher and more generalizable test of our argument. In all models, we use the same set of controls discussed earlier, while also holding constant economic hardship and household COVID-19 diagnoses.

Findings

We begin by examining the extent to which a sense of diffuse threat generated by anxiety-inducing experiences in 2020—measured by variables capturing economic hardship and household COVID-19 diagnoses—predicts gun buying. We also include the aforementioned

 $^{^{7} \}alpha = .69.$

 $^{^{8} \}alpha = .88.$

variables, ordered by the size of their effects.⁹ As expected, our threat variables are indeed important predictors of pandemic gun buying; as Figure 1 shows, both household COVID-19 and our economic hardship index are positive and statistically significant.¹⁰ Interestingly, we also find that, all else constant, parents were more likely to buy guns during the pandemic; this is consistent with our argument about threat, as pandemic-related disruptions to society may have impacted parents particularly strongly (given both their childcare needs and the financial costs of supporting a family). Finally, we find that several other factors that would theoretically be expected to predict gun buying are also significant, including Republican party identification. Taken together, these findings are consistent with the notion that the gun buying spike of 2020 was motivated by diffuse threat, leading—as we next show—to a more mistrustful, conspiracy-fearing population of gun-owners than before.

⁹ In Figure 1, we do not display pre-existing gun ownership since it is, in some sense, a lag of the outcome variable; it unsurprisingly had a large effect, as gun ownership predicts future gun ownership. We also do not display the coefficients for the region fixed effects, other party (since it is not clearly interpretable given that "other" is not well defined), or other race/ethnicity (since again it not clearly interpretable given that "other" is not well defined). See Appendix C for full regression tables.

¹⁰ To test whether these results were driven solely by pre-existing owners or new owners, we also ran regressions on each subsample. In each case, as would be expected, our threat variables are positive and significant. This indicates that threat has similar impacts among those Americans who did and did not own guns prior pandemic and is consistent with our contention that threat was an outsized motivator of pandemic gun buying. As a result of these similarities, we can be more confident in our claim that the *composition* of gun owners as a whole will now be more comprised of individuals motivated by threat, as it suggests pre-existing gun owners who chose *not* to buy more guns during the pandemic were, among the population of gun owners, those who were significantly less likely to feel a sense of threat.



Figure 1: Marginal Probability of Purchasing a Gun during the Pandemic

Estimates for threat variables (hardship index and COVID-19 in household) highlighted

Comparing the Views of Pre-Existing and New Gun Owners

We now turn to our primary hypotheses, which pertain to differences between the views of pre-existing and new gun owners. Our first hypothesis is that new gun owners will be more likely than pre-existing gun owners to hold conspiracy beliefs. These include, first, a belief that Trump was the true victor in the 2020 election, and, second, belief in conspiracy theories about the nature and effects of COVID-19 vaccines, such as whether they alter people's DNA or contain microchips (which we combine into an additive index).

Our findings are depicted in Figure 2, where we display three comparisons from models that can be found in full form in Appendix C. Specifically, we include results that compare preexisting gun owners to non-owners. These are important for interpreting the substantive meaning of both differences and similarities between new and pre-existing owners. As noted, we expect pre-existing owners to be more likely to hold conspiratorial views than the general public. We then present two tests of hypothesis 1, which posits that new gun buyers will be more likely to hold conspiracy beliefs than those who previously owned guns: one set of results includes a comparison against all pre-existing gun owners (perhaps the strictest test of our hypothesis) while the other focuses on new gun buyers relative to old gun buyers who did not buy during the pandemic. This latter comparison is interesting since those who did not buy at all were likely less motivated by threat and thus are even less likely to hold conspiratorial views. Each panel includes results for the 2020 election conspiracy, the vaccine conspiracy index, and each individual vaccine conspiracy item.

The first panel of Figure 2 shows that pre-existing gun owners, compared to all other respondents, are statistically significantly more likely to both believe that Trump won the 2020 election and hold conspiratorial views about COVID-19 vaccines. These findings are expected given prior work on the political views of gun owners (see, e.g., Joslyn 2020; Lacombe 2021). What about new gun owners? The second panel of Figure 2 shows partially consistent evidence with regard to the 2020 election: new gun owners are more likely to believe the conspiracy compared to pre-existing gun owners, although it does not reach conventional levels of statistical significance (p = 0.14). Moreover, given how strongly pre-existing gun ownership predicts a belief that Trump was the election's true victor (the first panel), a non-finding here is still notable as it suggests that first-time gun owners, rather than moderating the views of others in the gun owning community, are (at the very least) just as likely to believe the election was stolen.

In the case of vaccine views, we see (as predicted) that first-time gun buyers are significantly more likely to hold conspiracy beliefs than pre-existing gun owners, who themselves were already more likely than other respondents to hold such beliefs. This is the case for the scale as well as several of the individual items (with those that are not statistically significant at conventional levels nonetheless positive). This is clear support for hypothesis 1: the entry of new buyers into the ranks of gun ownership pulls an already conspiratorial group in a

more conspiratorial direction. This is accentuated by the last panel, where we see even stronger effects when comparing new gun buyers against pre-existing owners who did not buy during the pandemic.



We now turn to our second hypothesis, which is that new gun owners will be more likely to hold anti-system views than pre-existing gun owners. We measure such views through three different variables pertaining to institutional trust: Trust in government health institutions, trust in scientists, and trust in the news media. We present our findings in the same way we presented those pertaining to hypothesis 1, with the full results appearing in Appendix C. We again see, in the first panel of Figure 3, that pre-existing gun owners were less trusting of all three groups than other respondents; this coheres with the well-known political outlooks associated with gun ownership. More notable, however, is that the second panel strongly supports hypothesis 2: firsttime gun buyers report substantially less trust than pre-existing gun owners in all three cases; again, we see that these new gun owners pull an already low-trust group in an even less trusting direction. The final panel shows that relative to existing owners who did not buy during the pandemic, new owners also exhibit substantially less trust, marginally more so than with the comparison against all pre-existing owners.

Figure 3: Institutional trust among existing and new gun owners

Linear regression with controls (not shown). 90% and 95% uncertainty intervals use robust standard errors.



Taken all together, the findings presented in this section align with our argument. We examine five different relevant outcomes (as well as the component parts of the vaccine conspiracy index) and in four cases our findings confirm our hypotheses; in the fifth case— which pertains to the 2020 election—we find that new gun owners are no *less* likely than prior gun owners to hold the conspiratorial belief that Trump was the true victor and, indeed, that they are more likely to do so at the p < 0.14 significance level. Further, they are significantly more likely to hold that believe than pre-existing owners who did not buy during the pandemic.

Robustness Checks

We conducted two different types of checks to assess the robustness of our findings. The first further probes the stated motivations of gun buyers, focusing on the relationship between threat and first-time purchases. Recall that our earlier analyses looked at how anxiety-inducing events created diffuse feelings of threat that correlate with gun buying. The theoretical work on which we build makes clear that individuals do not necessarily need to consciously connect these diffuse feelings to explicitly-articulated rationales for buying guns. We can nonetheless look to such rationales for additional information, however, because our data include a question that asked those who purchased guns during the pandemic their reasons for doing so. The response

options include both hobbyist reasons—hunting and target shooting—and reasons that can be connected to threats, such as protection from crime.¹¹ Respondents could select all that apply. To be clear, this question was asked *only of those who bought guns during the pandemic*, which means that it excludes pre-existing gun owners who did *not* make additional purchases during or after March 2020; as a result, the sample used for analyses that include this question differs from the sample used in other parts of the paper and is less ideal for our purposes. Nonetheless, it provides some useful leverage about the reasons that people provided for buying guns.

Earlier, we showed that all 2020 gun buyers were more likely than the rest of the public to have faced threat-inducing events—namely, economic hardships and COVID-19 in the household. Yet, while elevated threats are associated with gun purchases by new and pre-existing owners alike, threats were more likely to be the stated rationale for purchase among first-time buyers. As Figure 4 shows, new gun owners attributed their purchases to threat-based motivations more frequently than pre-existing owners, while the latter were more likely to cite hunting and target-shooting, reasons that were already popular before the pandemic (see Parker et al. 2017). Thus, this influx of threat-driven buyers suggests that gun owners as a group are probably now more threat-driven than before the pandemic.¹² Note that, because we did not pose this question to pre-existing gun owners who did not purchase additional guns during the pandemic, the differences we identify between first-time and repeat gun buyers here very likely *understate* gaps between new and pre-existing owners.¹³ Although we expect (and, as discussed earlier, find) that all pandemic gun buyers were motivated by diffuse senses of threat, our finding

¹¹ "Hobbyist" response options consist of "hunting" and "target shooting." "Threat" reasons consist of: "Protection against crime," "protection against the government," "because of COVID-19," "because of lockdown and restrictions," "because of the election," and "protection against someone I know personally."

¹² They also may be quicker or more eager to use a gun to counter these threats, though other factors such as comfort and familiarity with using a firearm likely play a role.

¹³ Our logic here is that the group of pre-existing gun owners who did not make additional purchases during the pandemic would be expected to consist of hobbyists to a greater extent than the group of pandemic buyers.

here—which enables gun owners to articulate their reasons for buying—lends support to the notion that the composition of first-time gun buyers is unusually (consciously) motivated by threat.



Second, we also include a robustness check that pertains to our core hypotheses regarding the views of gun owners. Here, we explore the same dependent variables pertaining to conspiracy beliefs and trust discussed above, and, as independent variables, focus on our measures associated with diffuse threat. We do this by limiting the sample to all gun owners (i.e., first-time buyers, those who owned guns before the pandemic and bought more during it, and those who owned guns before the pandemic but did *not* buy more during it) and examining factors, described earlier, that would be linked with anxiety due to the pandemic. These are whether anyone in an individual's household was diagnosed with COVID-19 and whether an individual experienced pandemic-related financial hardships. Here, expect to find that household COVID-19 cases and economic hardship predict conspiracy beliefs and trust; our theory is that threat motivates both gun buying and the attitudes that we examine, which means our threat variables should predict our dependent variables. As Tables 1 and 2 show, this is indeed what we find: Both outcomes related to conspiracy beliefs are predicted by either household COVID-19 or the economic hardship index (or both), as are all three trust outcomes. In the latter case, we find that being a new gun owner is also significantly associated with reduced levels of trust (with new gun ownership positive but not significant in the other models); this is notable given that our theory does not necessitate a difference between new and repeat buyers here.

Table 1: Conspiracy Beliefs by Gun Ownership (Pre-Existing and New Owners)

| Dependent Variables: | Trump Won | Vaccine Conspiracy Index |
|----------------------|--------------|--------------------------|
| Variables | | |
| New Owner | 0.08 | 0.02 |
| | (0.07) | (0.01) |
| Hardship Index | 0.13 | 0.05^{**} |
| | (0.13) | (0.02) |
| Covid House | 0.30^{***} | 0.06^{***} |
| | (0.03) | (0.006) |
| Fit statistics | | |
| Observations | $7,\!424$ | 6,771 |
| Squared Correlation | 0.40483 | 0.13057 |
| $Pseudo R^2$ | 0.13319 | -0.67163 |
| BIC | $25,\!330.4$ | -2,102.1 |

Heteroskedasticity-robust standard-errors in parentheses Signif. Codes: ***: 0.01, **: 0.05, *: 0.1

| Dependent Variables: | Trust News Media | Trust Health Officials | Trust Scientists | |
|-----------------------|------------------|------------------------|------------------|--|
| Variables | | | | |
| New Owner | -0.14*** | -0.20*** | -0.10** | |
| | (0.05) | (0.04) | (0.05) | |
| Hardship Index | -0.13* | 0.02 | -0.02 | |
| | (0.08) | (0.07) | (0.07) | |
| Covid House | -0.03 | -0.07*** | -0.07*** | |
| | (0.02) | (0.02) | (0.02) | |
| Fit statistics | | | | |
| Observations | 7,398 | 7,385 | 7,405 | |
| Squared Correlation | 0.26650 | 0.30685 | 0.16678 | |
| Pseudo R ² | 0.11149 | 0.14012 | 0.07622 | |
| BIC | 18,530.5 | 16,868.4 | 16,633.9 | |

Table 2: Trust by Gun Ownership (Pre-Existing and New Owners)

 $Heterosked a sticity \hbox{-} robust\ standard \hbox{-} errors\ in\ parentheses$

Signif. Codes: ***: 0.01, **: 0.05, *: 0.1

These findings, in conjunction with our other findings, show that threat-based gun buyers differ from others, which indicates that the composition of the population of gun owners—as a result of the entry of a group of first-time gun buyers especially motivated by threat—has shifted during the pandemic, with a greater proportion holding conspiracy beliefs and reporting low levels of trust in important institutions.

Discussion and Conclusion

In this study, we have examined how the gun-buying surge of 2020—by bringing many millions of new Americans into the gun-owning community—may alter the group composition of gun owners moving forward. Gun owners are a notably important group, for at least two reasons. First, gun owners have been shown to be a crucial political constituency; participating in politics at unusually high rates, holding a distinct set of views, and comprising a key part of the Republican Party's electoral coalition, gun-owning Americans—led by the NRA—have played a large role not just in the realm of gun politics but in U.S. politics more broadly. As a result, potential shifts in their political behavior are important for both substantive and academic reasons. Second, gun owners are important because, by definition, they are armed; they possess the ability to—at least to some extent—address security concerns independently of the state and, indeed, to even potentially take on the state. As a result, their political attitudes and actions—particularly as it pertains to conspiracy beliefs and anti-system views—have important consequences for American democracy itself.

In this light, we believe our findings are troubling. We have demonstrated that the gun buying spike of 2020 was motivated in large part by threat, which prior work has shown to be associated with a distinct set of political views. We examine whether 2020's gun buyers—

particularly its *first-time* buyers who, by virtue of being new to the group, have the capacity to alter its composition—hold these views; we focus in particular on attitudes pertaining to the 2020 election and COVID-19 vaccines along with trust in public health, science, and media institutions. We find that new gun owners are in almost all cases more likely than pre-existing gun owners to hold conspiracy beliefs and anti-system views, even despite the fact that preexisting gun owners-relative to other Americans-are themselves more likely to hold such attitudes. In other words, our evidence contradicts the claims of some that first-time gun buyers will substantially moderate the sociopolitical meaning of gun ownership in the U.S. Rather, we find that the new gun owners of 2020 hold views that are *more* extreme than those of preexisting gun owners. Importantly, since new gun owners have beliefs that directionally echo those of prior owners (when viewed relative to the general population), they are unlikely to cause a fissure with the pre-existing population of gun owners, instead moving the group in a further conspiratorial and anti-system direction. Along these lines, given our finding that 2020 led to an increase in the number of people who have the means and motivation to act against the state and given events like the January 6, 2021 insurrection—subsequent work might extend our study by examining the relationship between gun buying and support for political violence.

Second, our findings demonstrate that when individuals take actions that stem from threat, there can be important downstream consequences that are not necessarily obvious. In this case, the events of 2020 caused a number of threats, which in turn motivated gun buying, which in turn has consequences for a number of different political outcomes. Understanding both how threat motivates actions and what sorts of consequences those actions have is thus important. More generally, threats typically do not prompt direct calculated actions to address their source. Instead, they often trigger a range of emotions that bias decision making. In the case of gun

buying, as mentioned, senses of diffuse threat can prompt gun purchases even when owning guns has no obvious connection to the threat. It also can alter reasoning as people seek attributions and explanations for the threat. In the case of the pandemic, these patterns seemed to connect with both gun buying and anti-system beliefs, a potentially very dangerous combination.

Finally, our work builds on important theories of group politics by demonstrating that disturbances—that is, social, political, or economic disruptions to the system—do not just encourage the mobilization of "potential groups" comprised of individuals who perceive their shared interests to be threatened, but can also lead to important changes in the *composition* of existing groups. In other words, when disturbances—such as a global pandemic—make individuals feel threatened, they may respond by entering the ranks of a pre-existing group. This decision may then have important consequences for who's part of that group and what sorts of views its members hold.

We have explored this pattern in the case of gun owners. A set of threatening conditions caused a surge in gun-buying, including among millions of individuals who did not previously own guns. This led to speculation that the relative diversity of these new gun owners—with some groups, such as Black Americans, buying guns for the first time at rates that slightly exceeded their pre-existing rates of gun ownership (while nonetheless comprising a relatively small proportion of pandemic purchasers)—would moderate the views of the gun-owning community. Our expectations—built on extant scholarship focused on the effects of disturbances and the threatening feelings they cause—were the opposite, however, and are borne out by our findings: rather than moderating the gun-owning community, the first-time gun buyers of 2020 have instead moved a group that was already especially likely to hold conspiracy beliefs and anti-system views in a more extreme direction. These findings—although, at this point, limited to one

group and one set of disturbances—suggest that subsequent work should consider not just how social, economic, and political disruptions mobilize groups, but also how they change the composition of groups that already exist. Such work could help explain how and why critical junctures caused by threatening events sometimes reorganize important lines of group-based political conflict in durable and otherwise unexpected ways.

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Appendix A: Sample Descriptive Features

Table A1 shows the demographics of our sample after reweighting to approximate the 2018 American Community Survey proportions for gender, age, race/ethnicity, educational attainment, urbanicity, and region. We find, in general, that our weights suffice for the other variables in the weighted sample to likewise reflect demographic patterns in the U.S. adult population.

| | All Respondents | Pre-Existing Gun Owners | Pandemic Purchasers | New Owners |
|-------------------------|-------------------------|-------------------------|------------------------|----------------------|
| Variable | N = 24,448 ¹ | N = 7,027 ¹ | N = 1,692 ⁷ | N = 318 ¹ |
| Gender | | | | |
| Female | 52% | 49% | 46% | 54% |
| Male | 48% | 51% | 54% | 46% |
| Race | | | | |
| White | 65% | 75% | 71% | 63% |
| Hispanic | 15% | 9.9% | 11% | 11% |
| Black | 12% | 9.4% | 11% | 17% |
| Asian American | 5.8% | 3.6% | 3.8% | 4.6% |
| Other | 2.5% | 2.3% | 3.2% | 4.3% |
| Community | | | | |
| rural | 14% | 20% | 19% | 15% |
| suburban | 55% | 56% | 54% | 55% |
| urban | 30% | 24% | 28% | 29% |
| Age | 47 | 48 | 41 | 40 |
| Education | | | | |
| High School or Less | 37% | 36% | 34% | 32% |
| Some College | 32% | 33% | 31% | 35% |
| Bachelor's Degree | 19% | 19% | 18% | 19% |
| Graduate Degree | 12% | 11% | 17% | 14% |
| Income Bracket | | | | |
| Less than \$25,000 | 25% | 17% | 16% | 19% |
| \$25,000 to \$74,999 | 44% | 46% | 40% | 49% |
| \$75,000 to \$149,999 | 24% | 29% | 33% | 24% |
| \$150,000 or more | 6.8% | 7.9% | 11% | 8.2% |
| Party | | | | |
| Republican | 25% | 37% | 37% | 26% |
| Democrat | 41% | 30% | 32% | 39% |
| Independent | 29% | 27% | 24% | 27% |
| Other | 5.9% | 5.9% | 6.5% | 7.9% |
| Ideology | | | | |
| Liberal | 33% | 25% | 29% | 32% |
| Moderate | 38% | 34% | 32% | 38% |
| Conservative | 29% | 41% | 38% | 31% |
| Trump won | 28% | 41% | 53% | 39% |
| Pre-Existing Owner | 29% | 100% | 81% | 0% |
| New Owner | 1.3% | 0% | 19% | 100% |
| ¹ Reweighted | | | | |

Table A1: Sample demographics.

Appendix B: Question Wording

Gun Ownership

Gun owner Do you or a member of your household own a gun?

OYes

ONo

Gun purchase Did you or a member of your household buy a gun during the COVID-19 pandemic (in the past 12 months)?

OYes

ONo

Pre-existing owner Did you or a member of your household own a gun before the COVID-19 pandemic (over 12 months ago)?

OYes

ONo

Reasons for Gun Buying

What were the reasons you or a member of your household decided to get a gun? (Please select all that apply)

| Hunting |
|-----------------------------------|
| Target shooting |
| Protection against crime |
| Protection against the government |

| Because of COVID-19 |
|--|
| Because of the lockdown and restrictions |
| Because of the election |
| Protection against someone I know personally |
| Other |

Conspiracy Beliefs

Respondents received a "1" for saying accurate and a "0" for saying inaccurate or not sure, except for the clinical trials item which was reverse coded (1 for inaccurate, 0 for accurate or not sure). The **vaccine conspiracy index** is the average of these scores.

Below are some statements about the COVID-19 vaccines that are currently being distributed. To the best of your knowledge, are those statements accurate or inaccurate?

| | Accurate | Inaccurate | Not sure |
|--|----------|------------|-------------|
| The COVID-19 vaccines will alter people's DNA. | 0 | 0 | 0 |
| The COVID-19 vaccines contain microchips that could track people. | 0 | 0 | 0 |
| The COVID-19 vaccines contain the lung tissue of aborted fetuses. | 0 | 0 | 0 |
| The COVID-19 vaccines can cause infertility, making it more difficult to get pregnant. | 0 | 0 | 0 |

| The COVID-19 vaccines were tested on thousands of people in clinical trials. | 0 | \cap | 0 |
|--|---|--------|---|
| | | U | U |
| | | | |

Institutional Trust

The **trust health officials** variable is an average of the following three items. These items were originally embedded in a longer list.

How much do you trust the following people and organizations to do the right thing to best handle the current coronavirus (COVID-19) outbreak?

| | A lot | Some | Not too much | Not at all |
|---|-------|------|-----------------|---------------|
| The FDA (Food and Drug Administration) | 0 | 0 | 0 | 0 |
| The CDC (Centers for Disease Control and Prevention) | 0 | 0 | 0 | 0 |
| Dr. Anthony Fauci of the National Institute of Health | 0 | 0 | Ο | 0 |

The *trust scientists* and *trust news media* variables appeared in the following matrix, which originally contained more items.

How much do you trust the following people and organizations to do the right thing to best handle the current coronavirus (COVID-19) outbreak?

| | A lot | Some | Not too much | Not at all |
|----------------------------|-------|------|--------------|------------|
| Scientists and researchers | | | | |
| Scientists and researchers | Ο | 0 | 0 | Ο |
| The name madia | | | | |
| The news media | 0 | 0 | 0 | 0 |
| | | | _ | |

COVID-19 Variables

The following two questions were used to construct the **COVID-19 House** variable. Respondents who answered "Yes" to the first question or any answer aside from "None" on the second question were coded as 1. All others were coded as 0.

Have you ever been diagnosed with coronavirus (COVID-19)?

O Yes, I was diagnosed by a medical professional

O No, I was not diagnosed but I think I may have it now

O No, I was not diagnosed but I think I had it previously and recovered

O No, I was not diagnosed and I do not think I ever had it

O I am not sure

How many members of your household (other than yourself) have been diagnosed with coronavirus (COVID-19)?

O None

One

O Two

O Three

O Four

O Five or more

The *economic hardship index* was constructed from the number of items a respondent checked in the following list.

Are you or someone in your household **currently experiencing** any of the following as a result of the coronavirus (COVID-19) outbreak? (Please select all that apply)

| Had to start working from home |
|---|
| Had their school or university closed |
| Had to stop or scale back work to take care of their children |
| Had to stop or scale back work to take care of someone who had COVID-19 |
| Had to take a cut in pay due to reduced hours or demand for their work |
| Was unable to make rent or mortgage payments |
| Was evicted from their home |
| Was laid off or lost a job |
| Was unable to make rent or mortgage payments |
| Was evicted from their home |
| None of the above |

Appendix C: Logit and Full Models

| Dependent Variable: | Bought Cun | During Pandomia |
|----------------------|----------------|---------------------|
| Model | | Logit |
| Model. | OL5 | Logit |
| Variables | | |
| (Intercept) | 0.0313^{***} | -3.766*** |
| | (0.0098) | (0.1782) |
| Pre-Existing Owner | 0.1748^{***} | 2.464^{***} |
| | (0.0048) | (0.0665) |
| Black | 0.0098* | 0.1712^{*} |
| | (0.0054) | (0.0995) |
| Asian | -0.0110* | -0.2024 |
| | (0.0066) | (0.1525) |
| Hispanic | -0.0084 | -0.1246 |
| mopume | (0.0061) | (0.1102) |
| Other Bace | 0.0301** | 0.5140*** |
| Other Maee | (0.0117) | (0.1578) |
| Famala | 0.0081** | 0.1999*** |
| remaie | -0.0081 | -0.1002 |
| | (0.0054) | (0.0509) |
| Children in HH | 0.0161 | 0.2816 |
| | (0.0040) | (0.0588) |
| Age (Decades) | -0.0121*** | -0.2263*** |
| <i>a</i> " | (0.0011) | (0.0184) |
| College | -0.0011 | -0.0120 |
| | (0.0036) | (0.0610) |
| HH Income $(10k)$ | 0.0018^{***} | 0.0290^{***} |
| | (0.0004) | (0.0055) |
| Rural | 0.0084 | 0.1340 |
| | (0.0057) | (0.0845) |
| Suburban | -0.0012 | -0.0157 |
| | (0.0037) | (0.0674) |
| White Evangelical | 0.0078^{*} | 0.1338^{**} |
| 0 | (0.0046) | (0.0647) |
| Democrat | 0.0072^{*} | 0.1229^{*} |
| | (0.0038) | (0.0747) |
| Republican | 0.0304*** | 0.4415*** |
| ree p dioffeeting | (0.0051) | (0.0769) |
| Other Party | 0.0226*** | 0.3728*** |
| o their i turty | (0.0079) | (0.1218) |
| Ideological Identity | -0.0002 | 0.0132 |
| Ideological Identity | (0.0002) | (0.0102) |
| Rogion: Roghing | 0.0102* | (0.0202) 0.2736* |
| Region. Rockies | (0.0192) | (0.1408) |
| Domiona Conthereost | (0.0110) | (0.1496) |
| Region: Southwest | (0.0075) | 0.3333 |
| | (0.0075) | (0.1290) |
| Region: Great Plains | -0.0035 | 0.0178 |
| | (0.0080) | (0.1394) |
| Region: Midwest | 0.0044 | 0.0888 |
| | (0.0061) | (0.1169) |
| region: South | 0.0118** | 0.2263** |
| | (0.0060) | (0.1106) |
| Region: Mid-Atlantic | -0.0025 | -0.0908 |
| | (0.0060) | (0.1303) |
| Region: New England | 0.0071 | 0.0846 |
| | (0.0068) | (0.1513) |
| Hardship Index | 0.0439^{***} | 0.6028^{***} |
| | (0.0140) | (0.1787) |
| COVID in HH | 0.0333^{***} | 0.4729^{***} |
| | (0.0039) | (0.0547) |
| Fit statistics | | · · · · · · |
| Decompetions | 91 0 49 | 24 049 |
| Conservations | 24,048 | 24,048 |
| Squarea Correlation | 0.12218 | 0.10277 |
| r seudo K- | 0.75332 | 0.21/34 |
| BIC | 1,298.6 | 10,325.0 |

Table C1: Pandemic Gun Buying by Pre-Existing Gun Ownership and Controls

| Dependent Variables: | Trump Won | Vaccine Conspiracy Index | Alters DNA | Contains Fetal Tissue | Contains Microchips | Causes Infertility | Not Tested |
|-----------------------|------------------------------|--------------------------|-----------------|-----------------------|---------------------|--------------------|-----------------|
| Model: | OLS | OLS | OLS | OLS | OLS | OLS | OLS |
| Variables | | | | | | | |
| (Intercept) | 1.421^{***} | 0.1658*** | 0.1012^{***} | 0.0482^{***} | 0.0886^{***} | 0.1084^{***} | 0.4893^{***} |
| | (0.0497) | (0.0083) | (0.0126) | (0.0121) | (0.0110) | (0.0126) | (0.0207) |
| Pre-Existing Owner | 0.1980^{***} | 0.0095*** | 0.0192^{***} | 0.0043 | 0.0075^{*} | 0.0231^{***} | -0.0058 |
| | (0.0193) | (0.0031) | (0.0046) | (0.0045) | (0.0041) | (0.0047) | (0.0073) |
| Black | -0.2036*** | 0.0391^{***} | 0.0243^{***} | 0.0259^{***} | 0.0327^{***} | 0.0325^{***} | 0.0816^{***} |
| | (0.0264) | (0.0047) | (0.0067) | (0.0068) | (0.0064) | (0.0071) | (0.0116) |
| Asian | 0.0599 | -0.0103* | -0.0047 | -0.0151^* | -0.0089 | -0.0218** | 0.0006 |
| | (0.0369) | (0.0060) | (0.0088) | (0.0083) | (0.0077) | (0.0087) | (0.0156) |
| Hispanic | -0.0793^{**} | 0.0144^{***} | 0.0114 | 0.0125 | 0.0105 | 0.0079 | 0.0321^{**} |
| | (0.0318) | (0.0055) | (0.0080) | (0.0079) | (0.0072) | (0.0084) | (0.0132) |
| Other Race | 0.0175 | 0.0266^{***} | 0.0259^{*} | 0.0019 | 0.0100 | 0.0133 | 0.0860^{***} |
| | (0.0527) | (0.0090) | (0.0141) | (0.0120) | (0.0114) | (0.0137) | (0.0223) |
| Female | -0.1210^{***} | -0.0235*** | -0.0425^{***} | -0.0357^{***} | -0.0438^{***} | -0.0243^{***} | 0.0290^{***} |
| | (0.0175) | (0.0028) | (0.0042) | (0.0041) | (0.0037) | (0.0042) | (0.0069) |
| Children in HH | 0.2544^{***} | 0.0469^{***} | 0.0429^{***} | 0.0526^{***} | 0.0523^{***} | 0.0476^{***} | 0.0407^{***} |
| | (0.0187) | (0.0032) | (0.0047) | (0.0046) | (0.0042) | (0.0050) | (0.0077) |
| Age (Decades) | -0.0075 | -0.0157*** | -0.0106*** | -0.0083*** | -0.0101*** | -0.0183*** | -0.0307*** |
| | (0.0050) | (0.0008) | (0.0012) | (0.0012) | (0.0011) | (0.0013) | (0.0022) |
| College | -0.1710^{***} | -0.0207*** | -0.0037 | 0.0033 | -0.0038 | -0.0030 | -0.0959*** |
| | (0.0176) | (0.0028) | (0.0040) | (0.0040) | (0.0034) | (0.0041) | (0.0073) |
| HH Income (10k) | -0.0055*** | -0.0006* | 0.0009^{**} | 0.0016^{***} | 0.0009^{**} | 0.0006 | -0.0069*** |
| | (0.0019) | (0.0003) | (0.0005) | (0.0004) | (0.0004) | (0.0005) | (0.0007) |
| Rural | 0.0812^{***} | 0.0055 | -0.0043 | -0.0031 | -0.0145^{**} | -0.0033 | 0.0525^{***} |
| | (0.0276) | (0.0047) | (0.0069) | (0.0068) | (0.0061) | (0.0071) | (0.0111) |
| Suburban | -0.0588^{***} | -0.0120*** | -0.0117^{**} | -0.0125^{***} | -0.0234^{***} | -0.0117^{**} | 0.0003 |
| | (0.0195) | (0.0032) | (0.0048) | (0.0048) | (0.0043) | (0.0048) | (0.0078) |
| White Evangelical | 0.3571^{***} | 0.0674^{***} | 0.0587^{***} | 0.0748^{***} | 0.0680^{***} | 0.0629^{***} | 0.0709^{***} |
| | (0.0230) | (0.0038) | (0.0057) | (0.0058) | (0.0052) | (0.0058) | (0.0086) |
| Democrat | -0.5571^{***} | -0.0067** | 0.0043 | 0.0213^{***} | 0.0156^{***} | 0.0042 | -0.0811^{***} |
| | (0.0208) | (0.0032) | (0.0048) | (0.0047) | (0.0043) | (0.0048) | (0.0085) |
| Republican | 1.103^{***} | 0.0266^{***} | 0.0309^{***} | 0.0313^{***} | 0.0168^{***} | 0.0422^{***} | 0.0122 |
| | (0.0274) | (0.0042) | (0.0062) | (0.0061) | (0.0054) | (0.0064) | (0.0099) |
| Other Party | 0.1308^{***} | 0.0200*** | 0.0153 | 0.0109 | -0.0003 | 0.0052 | 0.0687^{***} |
| | (0.0407) | (0.0063) | (0.0094) | (0.0085) | (0.0074) | (0.0095) | (0.0158) |
| Ideological Identity | 0.2133^{***} | 0.0101^{***} | 0.0045^{***} | 0.0065^{***} | 0.0010 | 0.0051^{***} | 0.0321^{***} |
| | (0.0067) | (0.0012) | (0.0017) | (0.0017) | (0.0016) | (0.0018) | (0.0025) |
| Region: Rockies | -0.0173 | -0.0249*** | -0.0260^{*} | -0.0182 | -0.0211^{*} | -0.0284^{**} | -0.0307 |
| | (0.0504) | (0.0096) | (0.0141) | (0.0134) | (0.0116) | (0.0138) | (0.0238) |
| Region: Southwest | 0.0313 | 0.0101 | -0.0017 | 0.0065 | 0.0108 | 0.0104 | 0.0177 |
| | (0.0378) | (0.0062) | (0.0093) | (0.0089) | (0.0082) | (0.0092) | (0.0150) |
| Region: Great Plains | -0.0141 | -0.0039 | -0.0178^{*} | -0.0181^{*} | -0.0055 | -0.0127 | 0.0294^{*} |
| | (0.0407) | (0.0071) | (0.0105) | (0.0099) | (0.0091) | (0.0103) | (0.0171) |
| Region: Midwest | 0.0041 | 0.0002 | -0.0103 | 0.0035 | 0.0007 | 0.0101 | -0.0059 |
| | (0.0323) | (0.0054) | (0.0080) | (0.0077) | (0.0070) | (0.0079) | (0.0130) |
| Region: South | 0.1421^{***} | 0.0006 | -0.0120 | 3.39×10^{-5} | 0.0005 | 0.0005 | 0.0121 |
| | (0.0313) | (0.0053) | (0.0079) | (0.0076) | (0.0069) | (0.0077) | (0.0126) |
| Region: Mid-Atlantic | 0.0204 | 0.0011 | -0.0166** | 0.0076 | 0.0004 | 0.0095 | 0.0002 |
| | (0.0339) | (0.0057) | (0.0084) | (0.0083) | (0.0076) | (0.0085) | (0.0139) |
| Region: New England | 0.0208 | -0.0124* | -0.0102 | -0.0059 | -0.0154^{**} | 0.0012 | -0.0373** |
| | (0.0385) | (0.0064) | (0.0094) | (0.0090) | (0.0075) | (0.0093) | (0.0163) |
| Hardship Index | -0.0129 | 0.0248^{**} | 0.0197 | 0.0459^{***} | 0.0440*** | 0.0575^{***} | -0.0399 |
| | (0.0641) | (0.0113) | (0.0170) | (0.0171) | (0.0159) | (0.0180) | (0.0246) |
| COVID in HH | 0.1878^{***} | 0.0432*** | 0.0508^{***} | 0.0493^{***} | 0.0424^{***} | 0.0524^{***} | 0.0202*** |
| | (0.0184) | (0.0032) | (0.0047) | (0.0047) | (0.0042) | (0.0049) | (0.0074) |
| Fit statistics | | | | | | | |
| Observations | 23,671 | 21,698 | 21,907 | 21,915 | 21,935 | 21,904 | 21,921 |
| Constant Constantion | 0.41203 | 0.10971 | 0.04745 | 0.05750 | 0.06619 | 0.05905 | 0.08126 |
| Squared Correlation | And the second second second | | | | | | |
| Pseudo R ² | 0.14071 | -0.34800 | 0.12753 | 0.15931 | 0.44176 | 0.13760 | 0.05921 |

| Table C | 2: | Conspiracy | Beliefs | by Gun | Ownership (| (Pre-Existing) | Owners vs. | Non-Owners |) |
|---------|----|------------|---------|--------|-------------|----------------|------------|------------|---|
|---------|----|------------|---------|--------|-------------|----------------|------------|------------|---|

| Dependent Variables: | Alters DNA | Contains Fetal Tissue | Contains Microchips | Causes Infertility | Not Tested |
|-----------------------|-----------------|-----------------------|-----------------------|--------------------|----------------------|
| Model: | Logit | Logit | Logit | Logit | Logit |
| Variables | | | | | |
| (Intercept) | -2.289^{***} | -2.979*** | -2.470*** | -2.170*** | -0.0448 |
| (1) | (0.1520) | (0.1574) | (0.1721) | (0.1478) | (0.0910) |
| Pre-Existing Owner | 0.2332*** | 0.0657 | 0.1311^{**} | 0.2611*** | -0.0238 |
| 0 | (0.0519) | (0.0538) | (0.0599) | (0.0508) | (0.0323) |
| Black | 0.3596*** | 0.4220*** | 0.6375^{***} | 0.4371^{***} | 0.3620*** |
| | (0.0857) | (0.0857) | (0.0931) | (0.0808) | (0.0506) |
| Asian | 0.0240 | -0.1035 | 0.0423 | -0.1796 | 0.0087 |
| | (0.1234) | (0.1335) | (0.1448) | (0.1258) | (0.0697) |
| Hispanic | 0.2034^{**} | 0.2536^{***} | 0.3128^{***} | 0.1635^{*} | 0.1424^{**} |
| | (0.0952) | (0.0968) | (0.1064) | (0.0927) | (0.0572) |
| Other Race | 0.3778^{**} | 0.0727 | 0.2962 | 0.2215 | 0.3742^{***} |
| | (0.1628) | (0.1939) | (0.2064) | (0.1653) | (0.0966) |
| Female | -0.4902^{***} | -0.3939^{***} | -0.6029*** | -0.2615^{***} | 0.1261^{***} |
| | (0.0486) | (0.0496) | (0.0551) | (0.0477) | (0.0308) |
| Children in HH | 0.4864^{***} | 0.5893^{***} | 0.7355^{***} | 0.5236^{***} | 0.1821^{***} |
| | (0.0518) | (0.0525) | (0.0586) | (0.0503) | (0.0336) |
| Age (Decades) | -0.1535^{***} | -0.1235^{***} | -0.2152^{***} | -0.2600^{***} | -0.1363^{***} |
| | (0.0166) | (0.0174) | (0.0200) | (0.0169) | (0.0096) |
| College | -0.0904^{*} | -0.0117 | -0.1447^{**} | -0.0822 | -0.4208^{***} |
| | (0.0532) | (0.0540) | (0.0594) | (0.0522) | (0.0323) |
| HH Income (10k) | 0.0100^{*} | 0.0168^{***} | 0.0096 | 0.0075 | -0.0312^{***} |
| | (0.0052) | (0.0051) | (0.0059) | (0.0051) | (0.0034) |
| Rural | -0.0243 | 0.0017 | -0.1372 | -0.0143 | 0.2232^{***} |
| | (0.0791) | (0.0804) | (0.0885) | (0.0778) | (0.0485) |
| Suburban | -0.1235^{**} | -0.1275^{**} | -0.2969^{***} | -0.1196^{**} | -0.0010 |
| | (0.0564) | (0.0570) | (0.0614) | (0.0551) | (0.0350) |
| White Evangelical | 0.6207^{***} | 0.7623^{***} | 0.8830^{***} | 0.6591^{***} | 0.3115^{***} |
| | (0.0584) | (0.0588) | (0.0674) | (0.0573) | (0.0374) |
| Democrat | 0.0185 | 0.2449^{***} | 0.2085^{***} | 0.0249 | -0.3591^{***} |
| | (0.0638) | (0.0653) | (0.0716) | (0.0617) | (0.0375) |
| Republican | 0.3789^{***} | 0.4405^{***} | 0.3407^{***} | 0.4886^{***} | 0.0460 |
| | (0.0717) | (0.0757) | (0.0862) | (0.0711) | (0.0422) |
| Other Party | 0.2081^{*} | 0.1644 | 0.0030 | 0.0917 | 0.2845^{***} |
| | (0.1174) | (0.1289) | (0.1504) | (0.1178) | (0.0680) |
| Ideological Identity | 0.0681*** | 0.0912*** | 0.0367* | 0.0727*** | 0.1435*** |
| D . D | (0.0192) | (0.0192) | (0.0217) | (0.0188) | (0.0111) |
| Region: Rockies | -0.2802 | -0.2077 | -0.3044 | -0.3106 | -0.1298 |
| | (0.1850) | (0.2007) | (0.2338) | (0.1946) | (0.1055) |
| Region: Southwest | -0.0208 | 0.0772 | 0.1577 | 0.1243 | 0.0823 |
| | (0.1078) | (0.1129) | (0.1225) | (0.1106) | (0.0671) |
| Region: Great Plains | -0.1869 | -0.2031 | -0.0481 | -0.1106 | 0.1318^{*} |
| D C MCL (| (0.1257) | (0.1345) | (0.1454) | (0.1290) | (0.0756) |
| Region: Midwest | -0.1104 | 0.0027 | 0.0394 | 0.1300 | -0.0224 |
| Device Couth | (0.0954) | (0.1006) | (0.1099) | (0.0976) | (0.0588) |
| Region: South | -0.1239 | 0.0227 | 0.0579 | (0.0042) | (0.0554) |
| Deniana Mid Atlantia | (0.0910) | (0.0907) | (0.1052) | (0.0947) | (0.0504) |
| Region: Mid-Atlantic | -0.2007 | 0.0904 | -0.0025 | 0.1189 | (0.0698) |
| Pogione Now England | (0.1020) | (0.1048) 0.1247 | (0.1157) 0.4847*** | (0.1055) | (0.0028) 0.1654** |
| Region: New England | -0.1400 | -0.1347 | -0.4647 | -0.0241 (0.1256) | -0.1034 (0.0754) |
| Handahin Indon | (0.1319) | (0.1422) | (0.1701) | (0.1550) | (0.0754) |
| maruship muex | 0.1760 | 0.4100 | (0.4501 | (0.1525) | -0.1/10 |
| COVID in HH | 0.1040) | 0.5262*** | 0.5401*** | 0.1000) | 0.1003) |
| | (0.0480) | (0.0202 | (0.0401 | (0.0200 | (0.0205) |
| | (0.0409) | (0.0002) | (0.0002) | (0.0400) | (0.0329) |
| Fit statistics | | | | | |
| Observations | 21,907 | 21,915 | 21,935 | 21,904 | 21,921 |
| Squared Correlation | 0.05174 | 0.06234 | 0.08047 | 0.06313 | 0.08129 |
| Pseudo R ² | 0.07377 | 0.08695 | 0.11949 | 0.09091 | 0.06186 |
| BIU | 12,980.8 | 12,714.6 | 10,443.7 | 13,338.3 | 28,363.9 |

Table C3: Conspiracy Beliefs by Gun Ownership (Pre-Existing Owners vs. Non-Owners)

| Dependent Variables: Model: | Trump Won OLS | Vaccine Conspiracy Index OLS | Alters DNA OLS | Contains Fetal Tissue OLS | Contains Microchips OLS | Causes Infertility OLS | Not Tested OLS |
|---|-----------------------|---------------------------------|--------------------|------------------------------|----------------------------|---------------------------|-------------------|
| Variables | | | | | | | |
| (Intercept) | 1 551*** | 0.9194*** | 0.1790*** | 0.0795*** | 0 1649*** | 0.1608*** | 0 5000*** |
| (intercept) | (0.1016) | (0.0170) | (0.0962) | (0.0242) | (0.0921) | (0.0969) | (0.0202) |
| Now Owner | (0.1010) | (0.0170) | 0.0203) | (0.0245) 0.0171 | (0.0231) | (0.0202) | (0.0392) |
| New Owner | (0.0710) | (0.0122) | (0.0304 | (0.0104) | (0.0102) | (0.0205) | (0.0317) |
| Dlasl. | (0.0719) 0.9740*** | (0.0133) | (0.0212) 0.0079 | (0.0194) | (0.0195) | (0.0205) 0.0122 | (0.0271) |
| DIACK | -0.2740 | (0.0223 | (0.0072 | 0.0194 | 0.0259 | (0.0123 | (0.0007) |
| A | (0.0580) | (0.0098) | (0.0140) | (0.0146) | (0.0140) | (0.0100) | (0.0227) |
| Asian | 0.0330 | -0.0334 | -0.0200 | -0.0241 | -0.0319 | -0.0428 | -0.0303 |
| TT: | (0.0962) | (0.0147) | (0.0221) | (0.0207) | (0.0186) | (0.0219) | (0.0349) |
| Hispanic | -0.1149 | 0.0172 | 0.0034 | 0.0304 | 0.0068 | -0.0129 | 0.0619 |
| | (0.0729) | (0.0134) | (0.0194) | (0.0191) | (0.0169) | (0.0198) | (0.0276) |
| Other Race | 0.2094** | 0.0110 | -0.0229 | -0.0170 | -0.0045 | -0.0084 | 0.1175*** |
| | (0.1038) | (0.0165) | (0.0240) | (0.0210) | (0.0215) | (0.0260) | (0.0412) |
| Female | -0.1937*** | -0.0347*** | -0.0635*** | -0.0492*** | -0.0572*** | -0.0373*** | 0.0351*** |
| CT 41 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 | (0.0331) | (0.0053) | (0.0082) | (0.0076) | (0.0068) | (0.0080) | (0.0122) |
| Children in HH | 0.3397*** | 0.0544*** | 0.0558*** | 0.0625*** | 0.0554*** | 0.0649*** | 0.0307** |
| | (0.0351) | (0.0061) | (0.0092) | (0.0087) | (0.0079) | (0.0097) | (0.0134) |
| Age (Decades) | -0.0070 | -0.0202*** | -0.0159^{***} | -0.0121*** | -0.0142^{***} | -0.0272^{***} | -0.0312^{***} |
| | (0.0097) | (0.0017) | (0.0024) | (0.0023) | (0.0021) | (0.0026) | (0.0039) |
| College | -0.1689^{***} | -0.0160*** | 0.0028 | 0.0093 | 0.0074 | -0.0039 | -0.0965^{***} |
| | (0.0329) | (0.0052) | (0.0077) | (0.0075) | (0.0062) | (0.0078) | (0.0129) |
| HH Income (10k) | -0.0088*** | -0.0013** | 0.0005 | 0.0008 | -0.0005 | 0.0002 | -0.0074^{***} |
| | (0.0034) | (0.0006) | (0.0008) | (0.0008) | (0.0007) | (0.0009) | (0.0013) |
| Rural | -0.0181 | -0.0067 | - 0.0256** | -0.0175 | -0.0311*** | 0.0027 | 0.0420^{**} |
| | (0.0505) | (0.0085) | (0.0129) | (0.0120) | (0.0111) | (0.0131) | (0.0188) |
| Suburban | -0.1375*** | -0.0295*** | -0.0342^{***} | -0.0254*** | -0.0436*** | -0.0180^{*} | -0.0211 |
| | (0.0421) | (0.0068) | (0.0105) | (0.0098) | (0.0091) | (0.0103) | (0.0152) |
| White Evangelical | 0.4338^{***} | 0.0656*** | 0.0684^{***} | 0.0756^{***} | 0.0669^{***} | 0.0655^{***} | 0.0492^{***} |
| | (0.0395) | (0.0064) | (0.0099) | (0.0095) | (0.0085) | (0.0102) | (0.0142) |
| Democrat | -0.5621^{***} | 0.0046 | 0.0109 | 0.0395^{***} | 0.0284^{***} | 0.0127 | -0.0759^{***} |
| | (0.0430) | (0.0067) | (0.0100) | (0.0096) | (0.0088) | (0.0103) | (0.0165) |
| Republican | 1.029^{***} | 0.0273*** | 0.0254^{**} | 0.0222** | 0.0272^{***} | 0.0281^{**} | 0.0351^{**} |
| | (0.0461) | (0.0071) | (0.0106) | (0.0101) | (0.0089) | (0.0112) | (0.0163) |
| Other Party | 0.1532^{**} | 0.0199^{*} | 0.0279 | 0.0070 | 0.0009 | -0.0106 | 0.0764^{***} |
| | (0.0772) | (0.0120) | (0.0188) | (0.0159) | (0.0142) | (0.0179) | (0.0283) |
| Ideological Identity | 0.2391*** | 0.0092*** | 0.0043 | 0.0070* [*] | -0.0039 | 0.0069^{**} | 0.0286*** |
| 0 | (0.0127) | (0.0023) | (0.0034) | (0.0034) | (0.0031) | (0.0035) | (0.0045) |
| Region: Rockies | -0.0942 | -0.0439*** | -0.0531** | -0.0223 | -0.0502*** | -0.0311 | -0.0724* |
| 0 | (0.0863) | (0.0160) | (0.0237) | (0.0215) | (0.0184) | (0.0241) | (0.0383) |
| Region: Southwest | 0.0726 | 0.0074 | -0.0081 | 0.0173 | -0.0123 | 0.0149 | 0.0203 |
| | (0.0735) | (0.0122) | (0.0189) | (0.0170) | (0.0159) | (0.0183) | (0.0274) |
| Region: Great Plains | -0.0676 | -0.0198 | -0.0609*** | -0.0148 | -0.0243 | -0.0238 | 0.0263 |
| | (0.0747) | (0.0129) | (0.0189) | (0.0178) | (0.0169) | (0.0191) | (0.0304) |
| Region: Midwest | 0.0047 | -0.0056 | -0.0181 | 0.0117 | -0.0095 | 0.0091 | -0.0200 |
| | (0.0654) | (0.0111) | (0.0169) | (0.0152) | (0.0146) | (0.0164) | (0.0247) |
| Region: South | 0.1784*** | -0.0035 | -0.0219 | 0.0072 | -0.0116 | 0.0076 | 0.0032 |
| | (0.0622) | (0.0107) | (0.0162) | (0.0146) | (0.0141) | (0.0158) | (0.0236) |
| Region: Mid-Atlantic | 0 2227*** | 0.0043 | -0.0258 | 0.0229 | -0.0002 | 0.0215 | -0.0048 |
| regioni ini ini ini | (0.0781) | (0.0129) | (0.0194) | (0.0180) | (0.0174) | (0.0194) | (0.0285) |
| Begion: New England | 0.0759 | -0.0114 | -0.0070 | 0.0082 | -0.0324* | 0.0116 | -0.0382 |
| Region. New England | (0.0886) | (0.0145) | (0.0232) | (0.0201) | (0.0167) | (0.0227) | (0.0352) |
| Hardship Index | 0.0753 | 0.0287 | 0.0565* | 0.0340 | 0.0687** | 0.0686* | -0.0784* |
| man and and and and and and and and and a | (0.1265) | (0.0218) | (0.0338) | (0.0326) | (0.0307) | (0.0353) | (0.0442) |
| COVID in HH | 0.2001*** | 0.0598*** | 0.0737*** | 0.0654*** | 0.0549*** | 0.0734*** | 0.0319** |
| 0.0 + 110 in 1111 | (0.0335) | (0.0058) | (0.0088) | (0.0084) | (0.0043 | (0.0091) | (0.0120) |
| | (0.0000) | (0.0030) | (0.0000) | (0.004) | (0.0014) | (0.0031) | (0.0123) |
| Fit statistics | | | | | | | |
| Observations | 7,575 | 6,905 | 6,975 | 6,969 | 6,986 | 6,966 | 6,986 |
| Squared Correlation | 0.39887 | 0.12802 | 0.07265 | 0.07242 | 0.08903 | 0.07985 | 0.08087 |
| Pseudo R ² | 0.13068 | -0.68542 | 0.13130 | 0.15749 | 0.33495 | 0.13121 | 0.05847 |
| BIC | 25,888.0 | -2,087.2 | 3,719.5 | 3,041.7 | 1,532.4 | 4,077.3 | 9,725.0 |
| | | | | | | | |

| Table C4: Conspira | cy Beliefs by Gur | ı Ownership (] | New Gun Owne | ers vs. Pre-Existin | g Gun Owners) |
|--------------------|-------------------|----------------|--------------|---------------------|---------------|

| Dopondont Variables: | Altors DNA | Containe Fotal Tissuo | Contains Microchine | Causes Infortility | Not Tostod |
|-----------------------|-----------------------|-----------------------|---------------------|--------------------|-----------------------|
| Model: | Logit | Logit | Logit | Logit | Logit |
| model: | Logit | Logit | LOGIU | Logit | Logit |
| Variables | | | | | |
| (Intercept) | -1.597*** | -2.715*** | -1.585*** | -1.611*** | 0.0910 |
| | (0.2529) | (0.2801) | (0.2990) | (0.2545) | (0.1708) |
| new_owner | 0.3429** | 0.1990 | 0.4203** | 0.1270 | 0.3555*** |
| D 1 1 | (0.1704) | (0.1808) | (0.1908) | (0.1678) | (0.1195) |
| Black | 0.1478 | 0.3346** | 0.4363*** | 0.1757 | 0.2477** |
| | (0.1506) | (0.1551) | (0.1642) | (0.1449) | (0.0994) |
| Asian | -0.1477 | -0.1466 | -0.2488 | -0.3013 | -0.1271 |
| TT: · | (0.2407) | (0.2777) | (0.3027) | (0.2446) | (0.1542) |
| Hispanic | (0.1789) | 0.4118 | 0.2228 | -0.0375 | 0.2008 |
| Out Print | (0.1782) | (0.1770) | (0.2014) | (0.1790) | (0.1190) |
| Other Race | -0.2859 | -0.2030 | -0.0248 | -0.0734 | 0.5055 |
| Famala | (0.3472) 0.6029*** | (0.3901) | (0.4150) | (0.3043) | (0.1817) 0.1525*** |
| Female | -0.0032 | -0.4832 (0.0826) | -0.0845 | -0.3202 | 0.1535 |
| Children in UU | 0.5144*** | (0.0830) | (0.0921) | 0.0762) | 0.1240** |
| | (0.0824) | (0.0882) | (0.0074) | (0.0804) | (0.0584) |
| Are (Decedee) | 0.1026*** | (0.0882) | 0.9694*** | 0.3130*** | 0.1350*** |
| Age (Decades) | (0.0272) | (0.0287) | (0.0336) | (0.0273) | (0.0171) |
| College | 0.0087 | 0.0287) | 0.0350) | 0.0213) | (0.0171) 0.4173*** |
| Conege | (0.0854) | (0.0002) | (0.0973) | (0.0837) | (0.0562) |
| HH Income $(10k)$ | 0.0018 | 0.0044 | -0.0133 | 0.0018 | -0.0325*** |
| IIII Income (Tok) | (0.0010) | (0.0087) | (0.0101) | (0.0081) | (0.0025) |
| Bural | -0.2066* | -0.1471 | -0 2994** | 0.0484 | 0.1778** |
| Iturui | (0.1243) | (0.1283) | (0.1406) | (0.1212) | (0.0822) |
| Suburban | -0 2961*** | -0 2316** | -0 4783*** | -0.1518 | -0.0938 |
| Suburban | (0.0968) | (0.0998) | (0.1067) | (0.0968) | (0.0674) |
| White Evangelical | 0.6177*** | 0.7431*** | 0.7966*** | 0.5998*** | 0.2128*** |
| | (0.0908) | (0.0956) | (0.1078) | (0.0895) | (0.0620) |
| Democrat | 0.0474 | 0.4078*** | 0.3452*** | 0.0814 | -0.3365*** |
| | (0.1113) | (0.1180) | (0.1318) | (0.1068) | (0.0729) |
| Republican | 0.3029*** | 0.3363*** | 0.4588^{***} | 0.3140^{***} | 0.1432^{**} |
| - | (0.1133) | (0.1254) | (0.1431) | (0.1127) | (0.0693) |
| Other Party | 0.3119^{*} | 0.1133 | 0.0476 | -0.0726 | 0.3208*** |
| | (0.1860) | (0.2203) | (0.2517) | (0.1941) | (0.1224) |
| Ideological Identity | 0.0577^{*} | 0.0951*** | -0.0170 | 0.0766** | 0.1273^{***} |
| | (0.0309) | (0.0335) | (0.0366) | (0.0308) | (0.0200) |
| Region: Rockies | -0.5394^{*} | -0.2823 | -0.8690** | -0.3241 | -0.3124^{*} |
| | (0.2891) | (0.3336) | (0.4014) | (0.2996) | (0.1711) |
| Region: Southwest | -0.0749 | 0.2085 | -0.1597 | 0.1410 | 0.0930 |
| | (0.1785) | (0.2013) | (0.2135) | (0.1922) | (0.1210) |
| Region: Great Plains | -0.6728^{***} | -0.1524 | -0.3127 | -0.2463 | 0.1163 |
| | (0.2224) | (0.2381) | (0.2495) | (0.2203) | (0.1334) |
| Region: Midwest | -0.1575 | 0.1653 | -0.0930 | 0.1039 | -0.0848 |
| | (0.1597) | (0.1850) | (0.1902) | (0.1743) | (0.1097) |
| Region: South | -0.1815 | 0.1321 | -0.0956 | 0.1082 | 0.0154 |
| | (0.1523) | (0.1777) | (0.1819) | (0.1681) | (0.1041) |
| Region: Mid-Atlantic | -0.2511 | 0.2379 | -0.0497 | 0.2075 | -0.0147 |
| | (0.1841) | (0.2008) | (0.2122) | (0.1915) | (0.1264) |
| Region: New England | -0.0493 | 0.0521 | -0.7868** | 0.1164 | -0.1697 |
| TT 11. T 1 | (0.2502) | (0.2907) | (0.3808) | (0.2594) | (0.1593) |
| Hardship Index | 0.4117 | 0.2397 | 0.5920** | 0.4537* | -0.3425* |
| | (0.2571) | (0.2767) | (0.2784) | (0.2491) | (0.1952) |
| COVID in HH | 0.0755*** | 0.0500*** | 0.0435*** | 0.6393*** | 0.1381^{**} |
| | (0.0809) | (0.0862) | (0.0948) | (0.0781) | (0.0561) |
| Fit statistics | | | | | |
| Observations | 6,975 | 6,969 | 6,986 | 6,966 | 6,986 |
| Squared Correlation | 0.08281 | 0.08621 | 0.11622 | 0.08436 | 0.08116 |
| Pseudo \mathbb{R}^2 | 0.09751 | 0.10201 | 0.14507 | 0.10757 | 0.06104 |
| BIC | 4,804.2 | 4,456.0 | 3,697.3 | 4,955.4 | 9,269.1 |

Table C5: Conspiracy Beliefs by Gun Ownership (New Gun Owners vs. Pre-Existing Gun Owners)

| Dependent Variables: Model: | Trump Won OLS | Vaccine Conspiracy Index OLS | Alters DNA OLS | Contains Fetal Tissue OLS | Contains Microchips OLS | Causes Infertility OLS | Not Tested OLS |
|--------------------------------|------------------|---------------------------------|-------------------|------------------------------|----------------------------|---------------------------|-------------------|
| Variables | | | | | | | |
| (Intercept) | 1.200^{***} | 0.1933*** | 0.1167^{***} | 0.0737^{***} | 0.1108^{***} | 0.1386^{***} | 0.5304^{***} |
| | (0.1069) | (0.0174) | (0.0268) | (0.0246) | (0.0227) | (0.0267) | (0.0435) |
| New Owner | 0.2867^{***} | 0.0633^{***} | 0.0708^{***} | 0.0481^{**} | 0.0688^{***} | 0.0496^{**} | 0.0843^{***} |
| | (0.0727) | (0.0132) | (0.0210) | (0.0194) | (0.0192) | (0.0204) | (0.0274) |
| Black | -0.2939^{***} | 0.0306*** | 0.0224 | 0.0375^{**} | 0.0180 | 0.0167 | 0.0606^{**} |
| | (0.0587) | (0.0100) | (0.0146) | (0.0152) | (0.0132) | (0.0151) | (0.0251) |
| Asian | 0.0790 | -0.0254* | -0.0328 | -0.0092 | -0.0216 | -0.0254 | -0.0206 |
| TT | (0.0940) | (0.0144) | (0.0201) | (0.0208) | (0.0176) | (0.0218) | (0.0383) |
| Hispanic | -0.1275* | 0.0035 | -0.0018 | -0.0068 | -0.0050 | -0.0141 | 0.0475 |
| Other Deer | (0.0731) | (0.0126) | (0.0191) | (0.0170) | (0.0157) | (0.0188) | (0.0310) |
| Other Race | (0.2028) | 0.0080 | -0.0321 | -0.0109 | -0.0354 (0.0152) | -0.0055 | (0.0464) |
| Formala | 0.0218 | 0.0040 | 0.0241) | (0.0227) | (0.0132) | 0.0273) | 0.0469*** |
| remaie | (0.0313) | -0.0040 | -0.0158 | -0.0199 | -0.0234 (0.0066) | -0.0037 | (0.0402) |
| Children in HH | 0.0347) | 0.0391*** | 0.0322*** | 0.0378*** | 0.0341*** | 0.0482*** | 0.0390*** |
| Children in fiff | (0.0379) | (0.0064) | (0.0096) | (0.0090) | (0.0080) | (0.0101) | (0.0550) |
| Age (Decades) | 0.0056 | -0.0185*** | -0.0145*** | -0.0110*** | -0.0101*** | -0.0223*** | -0.0345*** |
| nge (Decado) | (0.0102) | (0.0017) | (0.0025) | (0.0023) | (0.0020) | (0.0026) | (0.0043) |
| College | -0.2236*** | -0.0280*** | -0.0098 | -0.0043 | -0.0033 | -0.0108 | -0.1105*** |
| | (0.0345) | (0.0051) | (0.0076) | (0.0073) | (0.0057) | (0.0077) | (0.0141) |
| HH Income (10k) | -0.0118*** | -0.0023*** | -0.0009 | -0.0003 | -0.0010 | -0.0017** | -0.0074*** |
| · · · · | (0.0035) | (0.0006) | (0.0008) | (0.0008) | (0.0006) | (0.0008) | (0.0014) |
| Rural | 0.0684 | 0.0030 | -0.0120 | -0.0185 | -0.0131 | 0.0108 | 0.0543^{***} |
| | (0.0524) | (0.0083) | (0.0125) | (0.0118) | (0.0105) | (0.0128) | (0.0208) |
| Suburban | -0.0656 | -0.0191*** | -0.0128 | -0.0238** | -0.0259*** | -0.0124 | -0.0144 |
| | (0.0427) | (0.0067) | (0.0101) | (0.0097) | (0.0086) | (0.0099) | (0.0168) |
| White Evangelical | 0.2897^{***} | 0.0410^{***} | 0.0386^{***} | 0.0415^{***} | 0.0373^{***} | 0.0387^{***} | 0.0440^{***} |
| | (0.0427) | (0.0065) | (0.0100) | (0.0096) | (0.0085) | (0.0104) | (0.0160) |
| Democrat | -0.6493^{***} | -0.0172*** | -0.0102 | 0.0064 | 0.0024 | -0.0096 | -0.0794^{***} |
| | (0.0440) | (0.0066) | (0.0096) | (0.0092) | (0.0081) | (0.0101) | (0.0180) |
| Republican | 0.9861*** | 0.0146** | 0.0050 | 0.0132 | 0.0050 | 0.0055 | 0.0440** |
| | (0.0500) | (0.0073) | (0.0108) | (0.0102) | (0.0088) | (0.0114) | (0.0179) |
| Other Party | (0.0824) | 0.0057 | 0.0064 | -0.0033 | -0.0069 | -0.0304 | (0.0210) |
| Ideologiaal Identity | (0.0654) | (0.0121) | (0.0194) | (0.0100) | (0.0140) | (0.01777) | (0.0319) |
| ideological identity | (0.0136) | (0.0023) | (0.0130 | (0.0033) | (0.0032) | (0.0036) | (0.0204) |
| Region: Rockies | -0.0594 | -0.0536*** | -0.0597*** | -0.0319 | -0.0388** | -0.0453** | -0.0957** |
| Region: Roemes | (0.0975) | (0.0156) | (0.0224) | (0.0207) | (0.0180) | (0.0227) | (0.0418) |
| Region: Southwest | 0.0345 | -0.0031 | -0.0202 | -2.04×10^{-5} | -0.0221 | -0.0111 | 0.0339 |
| 0 | (0.0767) | (0.0121) | (0.0184) | (0.0165) | (0.0148) | (0.0178) | (0.0302) |
| Region: Great Plains | -0.0444 | -0.0199 | -0.0512*** | -0.0203 | -0.0221 | -0.0226 | 0.0226 |
| , e | (0.0785) | (0.0129) | (0.0187) | (0.0172) | (0.0161) | (0.0190) | (0.0334) |
| Region: Midwest | -0.0140 | -0.0088 | -0.0166 | 0.0033 | -0.0100 | -0.0033 | -0.0119 |
| | (0.0680) | (0.0112) | (0.0167) | (0.0149) | (0.0140) | (0.0162) | (0.0271) |
| Region: South | 0.1914^{***} | -0.0072 | -0.0259 | -0.0028 | -0.0141 | -0.0072 | 0.0206 |
| | (0.0654) | (0.0108) | (0.0162) | (0.0143) | (0.0136) | (0.0158) | (0.0259) |
| Region: Mid-Atlantic | 0.1841^{**} | -0.0027 | -0.0234 | 0.0005 | 5.06×10^{-5} | 0.0043 | 0.0023 |
| | (0.0805) | (0.0128) | (0.0191) | (0.0173) | (0.0166) | (0.0190) | (0.0312) |
| Region: New England | 0.0231 | -0.0284** | -0.0166 | -0.0097 | -0.0400*** | -0.0238 | -0.0478 |
| TT. 1.1.1. T. 1. | (0.0918) | (0.0138) | (0.0227) | (0.0192) | (0.0147) | (0.0208) | (0.0381) |
| Hardship Index | 0.0891 | -0.0039 | 0.0305 | (0.0227) | 0.0433 | 0.0257 | -0.1331 |
| COVID in HH | (0.1390) | (0.0225) | 0.0349) | 0.0327) | (0.0303) | 0.0370) | (0.0519) |
| | (0.0350) | (0.0059) | (0.0000) | (0.0086) | (0.0290) | (0.0400 | (0.0244) |
| | (0.0339) | (6.0009) | (0.0090) | (0.000) | (0.0014) | (0.0092) | (0.0144) |
| Fit statistics | | | | | | | |
| Observations | 6,113 | 5,601 | 5,657 | 5,651 | 5,663 | 5,653 | 5,665 |
| Squared Correlation | 0.45702 | 0.11211 | 0.03983 | 0.03669 | 0.04374 | 0.05353 | 0.09120 |
| rseudo K [*] | 0.15786 | -0.26983 | 0.13739 | 0.19708 | -0.46456 | 0.15621 | 0.06644 |
| ыс | 20,151.2 | -2,901.3 | 1,676.9 | 1,093.8 | -505.10 | 1,913.1 | 7,845.6 |

Table C6: Conspiracy Beliefs by Gun Ownership (New Gun Owners vs. Pre-Existing Gun Owners With No Pandemic Purchases)

Table C7: Conspiracy Beliefs by Gun Ownership (New Gun Owners vs. Pre-Existing Gun Owners With No Pandemic Purchases)

| Dependent Variables: Model: | Alters DNA Logit | Contains Fetal Tissue Logit | Contains Microchips Logit | Causes Infertility Logit | Not Teste Logit |
|--------------------------------|---------------------|--------------------------------|------------------------------|-----------------------------|--------------------|
| Variables | | | | | |
| (Intercept) | 2 001*** | 2 698*** | 1 9/1*** | 1.856*** | 0.1250 |
| (intercept) | (0.2208) | -2:030 | (0.2075) | -1.000 (0.2272) | (0.1200 |
| N | (0.5296) | (0.5500) | (0.3975) | (0.5272) | (0.1925) |
| New Owner | 0.6930**** | 0.5406 | 0.8811**** | 0.4939*** | 0.3713 |
| | (0.1681) | (0.1828) | (0.1878) | (0.1688) | (0.1223) |
| Black | 0.3170^{*} | 0.5540^{***} | 0.3520^{*} | 0.2399 | 0.2741^{**} |
| | (0.1816) | (0.1865) | (0.2125) | (0.1828) | (0.1117) |
| Asian | -0.3686 | -0.0314 | -0.2852 | -0.2067 | -0.0873 |
| | (0.3283) | (0.3196) | (0.3903) | (0.2913) | (0.1714) |
| Hispania | 0.0188 | 0.0244 | 0.0150 | 0.1065 | 0.2060 |
| mspanie | (0.0100 | -0.0244 | -0.0135 | (0.9927) | (0.1257 |
| | (0.2283) | (0.2484) | (0.2741) | (0.2237) | (0.1357 |
| Other Race | -0.5017 | -0.1648 | -1.061 | -0.0474 | 0.6008** |
| | (0.4615) | (0.4396) | (0.7531) | (0.3714) | (0.2084) |
| Female | -0.2053^{**} | -0.2736*** | -0.4426*** | -0.0610 | 0.2067^{**} |
| | (0.1027) | (0.1047) | (0.1189) | (0.1003) | (0.0603) |
| Children in HH | 0.4091*** | 0.5081*** | 0.6049*** | 0.5799*** | 0.1732** |
| | (0.1056) | (0.1103) | (0.1251) | (0, 1012) | (0.0666 |
| Area (Decedera) | 0.20023*** | 0.1759*** | 0.9251*** | 0.9119*** | 0.1598* |
| Age (Decades) | -0.2062 | -0.1752 | -0.2331 | -0.3113 | -0.1328 |
| ~ | (0.0332) | (0.0347) | (0.0410) | (0.0334) | (0.0192 |
| College | -0.1552 | -0.0917 | -0.1063 | -0.1729 | -0.4858^{*} |
| | (0.1072) | (0.1123) | (0.1194) | (0.1058) | (0.0627) |
| HH Income (10k) | -0.0092 | -0.0031 | -0.0221 | -0.0178 | -0.0333* |
| | (0.0110) | (0.0118) | (0.0138) | (0.0114) | (0.0066 |
| Rural | -0.1462 | -0.2344 | -0.1849 | 0.1265 | 0.2348* |
| | (0.1596) | (0.1606) | (0.1807) | (0.1515) | (0.0020 |
| Cuburban | 0.1600 | 0.2014** | 0.1007) | 0.1640 | 0.0652 |
| Bubuiban | -0.1090 | -0.3214 | -0.4467 | -0.1040 | -0.0055 |
| White Evangelical | (0.1275) | (0.1274) | (0.1456) | (0.1266) | (0.0757 |
| | 0.4915^{***} | 0.5736^{***} | 0.6649^{***} | 0.5049^{***} | 0.1917^{**} |
| | (0.1155) | (0.1223) | (0.1421) | (0.1155) | (0.0702) |
| Democrat | -0.2033 | 0.0565 | 0.0092 | -0.1666 | -0.3574^{**} |
| | (0.1405) | (0.1469) | (0.1609) | (0.1369) | (0.0808) |
| Republican | 0.0832 | 0.2278 | 0.1406 | 0.0865 | 0.1825^* |
| | (0.1341) | (0.1475) | (0.1735) | (0.1359) | (0.0771) |
| Other Denter | 0.0741 | 0.0259 | 0.1967 | 0.2711 | 0.9969* |
| Other Farty | 0.0741 | -0.0552 | -0.1207 | -0.5711 | 0.2808 |
| | (0.2258) | (0.2637) | (0.3067) | (0.2446) | (0.1390) |
| Ideological Identity | 0.1831^{***} | 0.1709^{***} | 0.0749 | 0.1732^{***} | 0.1202^{**} |
| | (0.0430) | (0.0466) | (0.0535) | (0.0446) | (0.0231) |
| Region: Rockies | -0.8575^{**} | -0.5363 | -0.8712^{*} | -0.6488^{*} | -0.4255^{*} |
| 0 | (0.3882) | (0.4223) | (0.5155) | (0.3862) | (0.1945) |
| Region: Southwest | -0.2540 | 0.0003 | -0.4509 | -0.1637 | 0 1573 |
| response outmost | (0.2904) | (0.2557) | (0.2045) | (0.2430) | (0.1954 |
| Doutons Cassi Dist | (0.2004) | (0.2001) | (0.4940) | (0.2439) | 0.1019 |
| negion: Great Plains | -0./310 | -0.3218 | -0.4363 | -0.3044 | 0.1019 |
| | (0.2719) | (0.2914) | (0.3178) | (0.2598) | (0.1490 |
| Region: Midwest | -0.1934 | 0.0620 | -0.1539 | -0.0313 | -0.0480 |
| | (0.1977) | (0.2251) | (0.2449) | (0.2102) | (0.1224) |
| Region: South | -0.3043 | -0.0137 | -0.2194 | -0.0780 | 0.0934 |
| - | (0.1900) | (0.2165) | (0.2340) | (0.2015) | (0.1165) |
| Region: Mid-Atlantic | -0.2689 | 0.0134 | -0.0148 | 0.0757 | 0.0109 |
| region. mid-relation | (0.2005) | (0.9547) | (0.9708) | (0.2240) | (0.1411 |
| Desta N. E.J. | (0.2020) | (0.2047) | (0.2708) | (0.2049) | (0.1411 |
| Region: New England | -0.2052 | -0.2108 | -1.443*** | -0.4023 | -0.2214 |
| | (0.3112) | (0.3722) | (0.6250) | (0.3441) | (0.1778) |
| Hardship Index | 0.4234 | 0.0291 | 0.5739 | 0.2823 | -0.5895^{*} |
| | (0.3588) | (0.3900) | (0.4002) | (0.3534) | (0.2335) |
| COVID in HH | 0.4340*** | 0.4878*** | 0.4996*** | 0.4567*** | 0.1070* |
| | (0.1021) | (0.1079) | (0.1215) | (0.0983) | (0.0638 |
| Eit statistics | (0.1021) | (0.2010) | (0.1210) | (0.0000) | 10.0000 |
| ru statistics | | | . | . | |
| Observations | 5,657 | 5,651 | 5,663 | 5,653 | 5,665 |
| Squared Correlation | 0.04213 | 0.03790 | 0.04931 | 0.05452 | 0.09172 |
| Pseudo R ² | 0.06850 | 0.06738 | 0.09744 | 0.09038 | 0.06941 |
| | | | | | |

| Dependent Variables: | Trust: News Media | Trust: Health Officials | Trust: Scientists |
|--------------------------------------|---------------------|-------------------------|---------------------|
| Model: | OLS | OLS | OLS |
| Variables | | | |
| (Intercept) | 2.326^{***} | 3.142^{***} | 3.506^{***} |
| | (0.0367) | (0.0312) | (0.0300) |
| Pre-Existing Owner | -0.1486^{***} | -0.0966^{***} | -0.0473^{***} |
| | (0.0129) | (0.0115) | (0.0109) |
| Black | 0.0668^{***} | -0.1062*** | -0.2406*** |
| | (0.0207) | (0.0176) | (0.0176) |
| Asian | 0.0941^{***} | 0.0634^{***} | -0.0449** |
| | (0.0286) | (0.0222) | (0.0223) |
| Hispanic | 0.0249 | -0.0324 | -0.0713^{***} |
| | (0.0235) | (0.0205) | (0.0200) |
| Other Race | -0.0253 | -0.1424^{***} | -0.1387^{***} |
| | (0.0400) | (0.0344) | (0.0361) |
| Female | -0.1079^{***} | -0.0125 | -0.0253^{**} |
| | (0.0123) | (0.0107) | (0.0102) |
| Children in HH | 0.1695^{***} | -0.0054 | -0.0527^{***} |
| | (0.0135) | (0.0117) | (0.0114) |
| Age (Decades) | 0.0624^{***} | 0.0486*** | 0.0322^{***} |
| ~ | (0.0038) | (0.0032) | (0.0031) |
| College | 0.0823^{***} | 0.0819^{***} | 0.1036^{***} |
| | (0.0128) | (0.0111) | (0.0104) |
| HH Income (10k) | 0.0056*** | 0.0092*** | 0.0120^{***} |
| | (0.0013) | (0.0011) | (0.0010) |
| Rural | -0.0674*** | -0.0716*** | -0.0672*** |
| ~ | (0.0193) | (0.0169) | (0.0164) |
| Suburban | -0.0756*** | -0.0366*** | -0.0129 |
| | (0.0139) | (0.0117) | (0.0112) |
| White Evangelical | 0.1363*** | -0.0163 | -0.0851*** |
| D | (0.0149) | (0.0132) | (0.0126) |
| Democrat | 0.3894*** | 0.3457^{***} | 0.1938*** |
| | (0.0152) | (0.0127) | (0.0121) |
| Republican | -0.1902*** | -0.1734*** | -0.0778*** |
| | (0.0180) | (0.0165) | (0.0153) |
| Other Party | -0.1856*** | -0.2118 | -0.1296*** |
| T11 | (0.0276) | (0.0270) | (0.0263) |
| Ideological Identity | -0.1060 | -0.1298 | -0.0924 |
| Denten Deulten | (0.0046) | (0.0039) | (0.0037) |
| Region: Rockies | -0.0464 | (0.0214) | (0.0817 |
| Dominu Couthmost | (0.0555) | (0.0314) | (0.0292) |
| Region: Southwest | -0.0597 | (0.0225) | (0.0222) |
| Pagion, Crost Plains | (0.0270) | (0.0255) | (0.0222) 0.0171 |
| negion. Great Flams | -0.0400 (0.0901) | (0.0472) | (0.0240) |
| Region: Midwest | -0.0291 | 0.02037 | (0.0249) 0.0119 |
| reston midwest | (0.0231) | (0.0198) | (0.0112) |
| Begion: South | 0.0232) | 0.0504*** | 0.0005 |
| reston. South | (0.0226) | (0.0192) | (0.0186) |
| Region: Mid-Atlantic | 0.0500** | 0.0562*** | 0.0292 |
| response and roundle | (0.0247) | (0.0207) | (0.0199) |
| Region: New England | 0.0602** | 0.0756*** | 0.0519** |
| The second second second | (0.0282) | (0.0240) | (0.0226) |
| Hardship Index | -0.0848* | 0.0765** | 0.0989*** |
| actual and and | (0.0447) | (0.0370) | (0.0367) |
| COVID in HH | -0.0112 | -0.0330*** | -0.0318*** |
| | (0.0130) | (0.0112) | (0.0109) |
| | (0.0100) | (0.0110) | (0.0100) |
| Fit statistics | 00.007 | 00 514 | 00.014 |
| Observations | 23,607 | 23,514 | 23,014 |
| Squared Correlation \mathbf{D}_{2} | 0.20277 | 0.23820 | 0.13882 |
| r seudo K ⁻ | 0.08145 | 0.10741 | U.U04U0 51 994 7 |
| | 00,598.9 | 05,401.4 | 01,834.7 |

Table C8: Trust by Gun Ownership (Pre-Existing Owners vs. Non-Owners)

| Dependent Variables: Model: | Trust: News Media OLS | Trust: Health Officials OLS | Trust: Scientists OLS |
|--------------------------------|--------------------------|--------------------------------|--------------------------|
| Variables | | | |
| (Intercept) | 2.305^{***} | 3.167^{***} | 3.449^{***} |
| (I) | (0.0684) | (0.0616) | (0.0586) |
| New Owner | -0.1596*** | -0.2198*** | -0.1327*** |
| | (0.0490) | (0.0456) | (0.0476) |
| Black | 0.0892^{**} | -0.1553*** | -0.3235*** |
| | (0.0422) | (0.0370) | (0.0368) |
| Asian | 0.0753 | 0.0370 | -0.0690 |
| | (0.0629) | (0.0556) | (0.0556) |
| Hispanic | 0.0689 | -0.0570 | -0.1252*** |
| P P P P | (0.0498) | (0.0479) | (0.0440) |
| Other Bace | 0.0935 | -0.2758*** | -0.1377** |
| | (0.0734) | (0.0650) | (0.0673) |
| Female | -0.0729*** | 0.0455** | 0.0177 |
| I cinaic | -0.0125 | (0.0196) | (0.0191) |
| Children in HH | 0.1581*** | 0.0205 | 0.0526** |
| | (0.0236) | (0.0293) | -0.0520 |
| Are (Decedes) | 0.0230/ | 0.0411) | 0.0200) |
| nge (Decades) | 0.0001 | 0.0491 | 0.0343 (0.0059) |
| Colloro | (U.UU07) 0.0622*** | (0.0039) 0.0715*** | (U.UUƏð) 0.1175*** |
| Conege | (0.0022^{-10}) | 0.0715 | (0.0107) |
| IIII I., | (0.0219) | (0.0200) | (0.0187) |
| HH Income (10k) | 0.0042* | 0.0112*** | 0.0140*** |
| D I | (0.0022) | (0.0020) | (0.0018) |
| Rural | -0.0701** | -0.0634** | -0.0704** |
| | (0.0326) | (0.0292) | (0.0285) |
| Suburban | -0.1059^{***} | -0.0581** | -0.0348 |
| | (0.0273) | (0.0239) | (0.0228) |
| White Evangelical | 0.1280^{***} | -0.0520** | -0.1418^{***} |
| | (0.0240) | (0.0222) | (0.0215) |
| Democrat | 0.4230^{***} | 0.3310^{***} | 0.1850^{***} |
| | (0.0297) | (0.0249) | (0.0233) |
| Republican | -0.1439^{***} | -0.1593^{***} | -0.0994^{***} |
| | (0.0291) | (0.0270) | (0.0255) |
| Other Party | -0.2927^{***} | -0.3281*** | -0.2178^{***} |
| | (0.0463) | (0.0493) | (0.0483) |
| Ideological Identity | -0.1284^{***} | -0.1576*** | -0.0969*** |
| | (0.0084) | (0.0072) | (0.0071) |
| Region: Rockies | -0.0780 | 0.0726 | 0.1363^{***} |
| _ | (0.0581) | (0.0533) | (0.0492) |
| Region: Southwest | -0.0567 | 0.0258 | 0.0487 |
| - | (0.0490) | (0.0454) | (0.0434) |
| Region: Great Plains | -0.1117** | 0.0278 | 0.0588 |
| 0 | (0.0509) | (0.0476) | (0.0460) |
| Region: Midwest | -0.0273 | 0.0501 | 0.0478 |
| J | (0.0437) | (0.0404) | (0.0392) |
| Region: South | -0.0065 | 0.0553 | 0.0617* |
| | (0.0418) | (0.0386) | (0.0374) |
| Region: Mid-Atlantic | 0.0328 | 0.0421 | 0.0460 |
| | (0.0514) | (0.0462) | (0.0446) |
| Region: New England | 0.0704 | 0.0508 | 0.0800 |
| reston. new England | (0.0576) | (0.0535) | (0.0509) |
| Hardshin Index | -0.0538 | 0.0000 | 0.0340 |
| maruship macx | -0.0000 | (0.1002) | (0.0540 |
| COVID in UU | (0.0190) | 0.0704) | 0.0701) |
| | -0.0192 | -0.0025 | -0.0002^{-11} |
| | (0.0220) | (0.0201) | (0.0194) |
| Fit statistics | | | |
| Observations | 7,548 | 7,534 | 7,555 |
| Conservation of Conservations | 0.21736 | 0.24887 | 0.14257 |
| Squared Correlation | | | |
| Pseudo \mathbb{R}^2 | 0.08810 | 0.10913 | 0.06388 |

Table C9: Trust by Gun Ownership (New Owners vs. Pre-Existing Owners)

| Dependent Variables: Model: | Trust: News Media OLS | Trust: Health Officials OLS | Trust: Scientists OLS |
|--------------------------------|--------------------------|--------------------------------|--------------------------|
| Variables | | | |
| (Intercept) | 2.302^{***} | 3.199^{***} | 3.476^{***} |
| | (0.0764) | (0.0675) | (0.0639) |
| New Owner | -0.1343*** | -0.2448*** | -0.1586*** |
| | (0.0497) | (0.0461) | (0.0478) |
| Black | 0.1228^{***} | -0.1357^{***} | -0.3245^{***} |
| | (0.0462) | (0.0401) | (0.0400) |
| Asian | 0.0724 | 0.0099 | -0.0502 |
| | (0.0667) | (0.0596) | (0.0589) |
| Hispanic | 0.0827 | -0.0155 | -0.1164** |
| - | (0.0554) | (0.0526) | (0.0491) |
| Other Race | 0.1680^{*} | -0.2102*** | -0.0466 |
| | (0.0859) | (0.0733) | (0.0723) |
| Female | -0.0509** | 0.0290 | -0.0171 |
| | (0.0238) | (0.0215) | (0.0205) |
| Children in HH | 0.1250*** | -0.0442* | -0.0645*** |
| | (0.0266) | (0.0237) | (0.0228) |
| Age (Decades) | 0.0677*** | 0.0523*** | 0.0333*** |
| | (0.0074) | (0.0065) | (0.0062) |
| College | 0.0660*** | 0.0967*** | 0 1317*** |
| 0011080 | (0.0242) | (0.0218) | (0.0203) |
| HH Income (10k) | 0.0021 | 0.0102*** | 0.0135*** |
| IIII Income (Iok) | (0.0021) | (0.0022) | (0.0020) |
| Bural | -0.0299 | -0.0437 | -0.0618** |
| Iturai | (0.0255) | (0.0324) | (0.0311) |
| Suburban | 0.0752** | 0.0363 | 0.0205 |
| Suburban | (0.0302) | (0.0303) | (0.0200) |
| White Evengelies | (0.0302) | 0.0558** | (0.0249) 0.1452*** |
| white Evangencai | (0.0003) | -0.0338 | -0.1455 |
| Domograf | 0.2604*** | 0.2102*** | 0.1876*** |
| Democrat | (0.0222) | (0.0270) | (0.0240) |
| Donublicon | (0.0522) 0.1654*** | (0.0270) 0.1641*** | (0.0249) 0.0014*** |
| Republican | -0.1034 | -0.1041 | -0.0914 |
| Othen Dentry | (0.0520) | (0.0298) | (0.0277) 0.2041*** |
| Other Farty | -0.3132 | -0.2944 | -0.2041 |
| Idealamical Identity | (0.0010) 0.1961*** | (0.0558) 0.1614*** | (0.0556) |
| ideological identity | -0.1201 | -0.1014 | -0.0970 |
| Deuten Deulte | (0.0094) 0.1949* | (0.0085) | (0.0060) |
| Region: Rockies | -0.1242° | 0.0758 | 0.1085 |
| Denten Conthered | (0.0043) | (0.0598) | (0.0520) |
| Region: Southwest | -0.0950 | 0.0204 | 0.0399 |
| Denten Court Distan | (0.0540) | (0.0490) | (0.0405) |
| Region: Great Plains | -0.1160° | 0.0144 | 0.0513 |
| Dominant Mill and | (0.0559) | (0.0516) | (0.0490) |
| Region: Midwest | -0.0504 | 0.0340 | 0.0414 |
| | (0.0478) | (0.0434) | (0.0411) |
| Region: South | -0.0591 | 0.0203 | 0.0542 |
| | (0.0460) | (0.0418) | (0.0393) |
| Region: Mid-Atlantic | 0.0167 | 0.0546 | 0.0722 |
| | (0.0505) | (0.0499) | (0.0468) |
| Region: New England | 0.0680 | 0.0772 | 0.1267^{**} |
| TT 11. T 1 | (0.0636) | (0.0567) | (0.0519) |
| Hardship Index | -0.1956** | 0.1194 | 0.0670 |
| | (0.0955) | (0.0803) | (0.0795) |
| COVID in HH | -0.0280 | -0.0553** | -0.0475** |
| | (0.0247) | (0.0225) | (0.0212) |
| Fit statistics | | | |
| Observations | 6,094 | 6.079 | 6.095 |
| Squared Correlation | 0.19127 | 0.24939 | 0.14890 |
| Pseudo \mathbb{R}^2 | 0.07755 | 0.11006 | 0.06806 |
| DIC | 15 622 4 | 14 336 4 | 13 691 0 |

Table C10: Trust by Gun Ownership (New Owners vs. Pre-Existing Owners With No Pandemic Purchases)

| Dependent Variables: | Hunting | Target Shooting | Crime | Government | COVID | Lockdowns | Election | Protection | Other |
|-----------------------|------------|-----------------|----------------|-----------------|-----------------|-----------------------|-----------------|-----------------|---------------|
| Model: | OLS | OLS | OLS | OLS | OLS | OLS | OLS | OLS | OLS |
| Variables | | | | | | | | | |
| (Intercept) | 0.4688*** | 0 5861*** | 0.5000*** | 0.3284*** | 0.1140** | 0.0586 | 0.1257** | 0.1361*** | 0.1748*** |
| (intercept) | (0.0601) | (0.0755) | (0.0720) | (0.0500) | (0.0482) | (0.0520) | (0.0584) | (0.0248) | (0.0510) |
| Now Owner | 0.9186*** | 0.1725*** | 0.1108*** | 0.0330) | 0.0403) | 0.0492* | 0.0045 | 0.0348) | 0.0170 |
| New Owner | -0.2160 | -0.1735 | (0.0255) | -0.0341 | (0.0211) | (0.0423 | -0.0045 | -0.0180 | -0.0170 |
| DI I | (0.0210) | (0.0272) | (0.0255) | (0.0215) | (0.0191) | (0.0227) | (0.0218) | (0.0127) | (0.0188) |
| Васк | -0.1571 | -0.1044 | 0.0230 | -0.0186 | 0.0652 | 0.0142 | 0.0178 | 0.0315 | 0.0238 |
| | (0.0343) | (0.0394) | (0.0391) | (0.0327) | (0.0312) | (0.0312) | (0.0337) | (0.0243) | (0.0291) |
| Asian | -0.1217** | 0.0369 | 0.0108 | -0.0822* | -0.0073 | -0.0839** | -0.0777* | -0.0504*** | 0.0209 |
| | (0.0583) | (0.0699) | (0.0598) | (0.0437) | (0.0404) | (0.0349) | (0.0398) | (0.0105) | (0.0452) |
| Hispanic | -0.0702 | -0.1240^{**} | -0.0302 | -0.0208 | 0.0537 | 0.0247 | -0.0679** | 0.0172 | 0.0260 |
| | (0.0447) | (0.0491) | (0.0519) | (0.0396) | (0.0357) | (0.0389) | (0.0326) | (0.0265) | (0.0375) |
| Other Race | 0.0701 | 0.0814 | 0.0463 | -0.0162 | 0.0262 | 0.0390 | 0.0523 | 0.0488 | 0.0287 |
| | (0.0601) | (0.0667) | (0.0572) | (0.0498) | (0.0415) | (0.0537) | (0.0543) | (0.0382) | (0.0515) |
| Female | -0.0726*** | -0.0962^{***} | 0.0447^{*} | -0.1225^{***} | -0.0664^{***} | -0.0525*** | -0.0522^{***} | -0.0022 | 0.0377^{**} |
| | (0.0232) | (0.0253) | (0.0235) | (0.0206) | (0.0162) | (0.0193) | (0.0199) | (0.0113) | (0.0164) |
| Children in HH | 0.0326 | 0.0082 | -0.0246 | -0.0226 | 0.0230 | 0.0115 | -0.0282 | 0.0062 | -0.0054 |
| | (0.0227) | (0.0249) | (0.0229) | (0.0195) | (0.0148) | (0.0184) | (0.0180) | (0.0117) | (0.0170) |
| Age (Decades) | -0.0388*** | -0.0025 | 0.0061 | -0.0222*** | -0.0069 | -0.0034 | -0.0039 | -0.0114^{***} | 0.0092^{*} |
| | (0.0073) | (0.0082) | (0.0076) | (0.0065) | (0.0044) | (0.0057) | (0.0064) | (0.0036) | (0.0056) |
| College | -0.0009 | -0.0359 | -0.0544** | 0.0089 | 0.0249 | 0.0229 | 0.0093 | 0.0130 | -0.0095 |
| 0 | (0.0241) | (0.0265) | (0.0248) | (0.0208) | (0.0155) | (0.0196) | (0.0202) | (0.0112) | (0.0183) |
| HH Income (10k) | -0.0007 | 0.0003 | -0.0041* | -0.0033* | 0.0024 | 7.26×10^{-5} | -0.0003 | -0.0016 | 0.0011 |
| | (0.0022) | (0.0024) | (0.0023) | (0.0019) | (0.0016) | (0.0018) | (0.0019) | (0.0011) | (0.0017) |
| Bural | 0.1012*** | 0.0349 | -0.0384 | 0.0247 | -0.0160 | 0.0036 | -0.0095 | -0.0275 | 0.0227 |
| - our ou | (0.0351) | (0.0370) | (0.0345) | (0.0278) | (0.0241) | (0.0274) | (0.0271) | (0.0181) | (0.0247) |
| Suburban | 0.0015 | 0.0228 | -0.0470* | 0.0487** | -0.0159 | 0.0116 | 0.0154 | -0.0264* | 0.0168 |
| Suburban | (0.0013) | (0.0228) | (0.0283) | (0.0407 | (0.0210) | (0.0228) | (0.0231) | (0.0204) | (0.0180) |
| White Evengelicel | 0.0600*** | 0.0207 | 0.0526** | 0.0232) | 0.0657*** | 0.0223) | 0.0267* | 0.0260** | 0.0256** |
| winte Evangencar | (0.0099) | -0.0201 | -0.0550 | -0.0019 | (0.0037) | (0.0214) | -0.0307 | (0.0200) | -0.0350 |
| Domocrat | 0.0207) | 0.0112 | 0.0007 | 0.0222) | 0.0072 | 0.0211) | 0.0210) | 0.0250** | 0.0002*** |
| Democrat | (0.0714) | (0.0245) | (0.0220) | -0.0199 | (0.0013 | -0.0239 | (0.0322) | -0.0350 | -0.0993 |
| Danahliana | (0.0511) | (0.0545) | (0.0529) | (0.0240) | (0.0227) | (0.0244) | (0.0220) | (0.0172) | (0.0232) |
| Republican | 0.0048 | 0.0051 | -0.0141 | 0.0040 | -0.0138 | 0.0524 | (0.0900 | -0.0202 | -0.0071 |
| OUL D I | (0.0301) | (0.0332) | (0.0306) | (0.0263) | (0.0206) | (0.0247) | (0.0245) | (0.0149) | (0.0232) |
| Other Party | -0.0547 | -0.0107 | 0.0526 | 0.0713 | 0.0140 | 0.1213 | 0.0278 | -0.0093 | 0.0830 |
| T1 1 1 1 1 1 | (0.0436) | (0.0521) | (0.0459) | (0.0432) | (0.0319) | (0.0441) | (0.0368) | (0.0276) | (0.0462) |
| Ideological Identity | 0.0070 | -0.0011 | 0.0433^{***} | 0.0180*** | -0.0003 | 0.0088 | 0.0150** | -0.0042 | -0.0108** |
| | (0.0076) | (0.0083) | (0.0081) | (0.0066) | (0.0062) | (0.0065) | (0.0063) | (0.0038) | (0.0050) |
| Region: Rockies | 0.0772 | -0.0324 | 0.0703 | -0.1459*** | -0.0401 | 0.0009 | -0.0455 | 0.0022 | 0.0002 |
| | (0.0611) | (0.0657) | (0.0606) | (0.0490) | (0.0375) | (0.0428) | (0.0475) | (0.0249) | (0.0451) |
| Region: Southwest | -0.0780 | -0.0012 | 0.0651 | -0.0457 | -0.0098 | 0.0455 | 0.0074 | -0.0118 | 0.0273 |
| | (0.0508) | (0.0573) | (0.0529) | (0.0468) | (0.0386) | (0.0396) | (0.0444) | (0.0224) | (0.0398) |
| Region: Great Plains | -0.0127 | -0.0870 | 0.0986^{*} | -0.0365 | -0.0410 | 0.0491 | -0.0448 | 0.0104 | -0.0259 |
| | (0.0583) | (0.0619) | (0.0577) | (0.0535) | (0.0377) | (0.0443) | (0.0463) | (0.0260) | (0.0413) |
| Region: Midwest | -0.0063 | -0.0316 | 0.0292 | -0.0557 | -0.0125 | 0.0139 | -0.0072 | 0.0135 | -0.0170 |
| | (0.0487) | (0.0521) | (0.0486) | (0.0438) | (0.0357) | (0.0356) | (0.0407) | (0.0220) | (0.0344) |
| Region: South | -0.0135 | -0.1109^{**} | 0.0631 | -0.0959^{**} | -0.0183 | 0.0215 | -0.0250 | 0.0298 | -0.0346 |
| | (0.0458) | (0.0488) | (0.0454) | (0.0410) | (0.0334) | (0.0332) | (0.0382) | (0.0213) | (0.0324) |
| Region: Mid-Atlantic | -0.0116 | -0.0391 | 0.0428 | -0.0643 | -0.0161 | 0.0339 | -0.0208 | -0.0003 | -0.0507 |
| | (0.0547) | (0.0590) | (0.0559) | (0.0488) | (0.0409) | (0.0413) | (0.0455) | (0.0240) | (0.0359) |
| Region: New England | -0.0650 | -0.0145 | 0.0743 | -0.1168** | 0.0060 | 0.0439 | -0.0310 | 0.0119 | -0.0371 |
| 0 0 | (0.0611) | (0.0681) | (0.0647) | (0.0506) | (0.0443) | (0.0473) | (0.0501) | (0.0289) | (0.0445) |
| Hardship Index | 0.1003 | 0.0949 | 0.1020 | 0.0814 | 0.0935^{*} | 0.1584^{**} | 0.0511 | 0.0934^{**} | -0.0066 |
| L | (0.0684) | (0.0721) | (0.0647) | (0.0606) | (0.0552) | (0.0615) | (0.0574) | (0.0472) | (0.0496) |
| COVID in HH | -0.0228 | -0.0238 | -0.0695*** | -0.0032 | 0.0236 | 0.0168 | 0.0103 | -0.0132 | -0.0179 |
| | (0.0224) | (0.0245) | (0.0229) | (0.0193) | (0.0148) | (0.0184) | (0.0190) | (0.0112) | (0.0169) |
| | (0.0001) | (0.0210) | (0.0000) | (0.0100) | (0.0110) | (0.0101) | (0.0100) | (0.0112) | (0.0100) |
| Fit statistics | | | | | | | | | |
| Observations | 1,810 | 1,810 | 1,810 | 1,810 | 1,810 | 1,810 | 1,810 | 1,810 | 1,810 |
| Squared Correlation | 0.10673 | 0.05535 | 0.09040 | 0.05947 | 0.06467 | 0.03107 | 0.03197 | 0.03589 | 0.04774 |
| Pseudo R ² | 0.08653 | 0.04025 | 0.07168 | 0.06569 | 0.12989 | 0.04042 | 0.03954 | -0.22545 | 0.09047 |
| BIC | 2,359.2 | 2,659.9 | 2,423.5 | 1,780.9 | 1,013.2 | 1,558.8 | $1,\!631.4$ | -157.10 | 1,092.7 |

Table C11: Reasons for Pandemic Gun Purchases (New Owners vs. Pre-Existing Owners)

| Dependent Variables: | Hunting | Target Shooting | Crime | Government | COVID | Lockdowns | Election | Protection | Other |
|-----------------------|-----------------|-----------------------|----------------|-----------------------|-----------------|----------------------|-----------------|--------------------|---------------------|
| Model: | Logit | Logit | Logit | Logit | Logit | Logit | Logit | Logit | Logit |
| Variables | | | | | | | | | |
| (Intercept) | 0.0650 | 0.4231 | -0.0129 | -0.5713 | -2.178^{***} | -2.528^{***} | -1.993^{***} | -1.104 | -1.640^{***} |
| | (0.3586) | (0.3293) | (0.3469) | (0.3972) | (0.4943) | (0.4859) | (0.4627) | (0.6872) | (0.5017) |
| New Owner | -1.460^{***} | -0.8256*** | 0.6478^{***} | -0.2943 | 0.2493 | 0.3312^{*} | -0.0343 | -0.3837 | -0.1528 |
| | (0.1898) | (0.1409) | (0.1510) | (0.1809) | (0.2020) | (0.1693) | (0.1760) | (0.3174) | (0.2000) |
| Black | -1.022^{***} | -0.8201*** | 0.1031 | -0.1699 | 0.7610^{***} | 0.1051 | 0.1540 | 0.5108 | 0.2756 |
| | (0.2506) | (0.2127) | (0.1993) | (0.2690) | (0.2757) | (0.2615) | (0.2507) | (0.3738) | (0.2891) |
| Asian | -0.7083^{*} | 0.1579 | 0.0267 | -0.7509 | 0.0551 | -1.058^{*} | -0.7374 | -14.57^{***} | 0.2172 |
| | (0.3888) | (0.2998) | (0.2931) | (0.4667) | (0.5167) | (0.6271) | (0.4813) | (0.2487) | (0.3929) |
| Hispanic | -0.3973 | -0.5784^{**} | -0.1622 | -0.1554 | 0.6648^{**} | 0.2079 | -0.6146^{*} | 0.3493 | 0.2545 |
| | (0.2597) | (0.2403) | (0.2494) | (0.2946) | (0.3254) | (0.2869) | (0.3571) | (0.4128) | (0.3122) |
| Other Race | 0.3724 | 0.3529 | 0.2668 | -0.1253 | 0.3137 | 0.2737 | 0.3555 | 0.7564 | 0.2299 |
| | (0.2898) | (0.2785) | (0.3481) | (0.3584) | (0.5094) | (0.3655) | (0.3429) | (0.5020) | (0.3866) |
| Female | -0.3955^{***} | -0.4242^{***} | 0.1951^{*} | -0.8557^{***} | -0.6931^{***} | -0.4227^{***} | -0.3860^{***} | -0.0599 | 0.4251^{**} |
| | (0.1217) | (0.1105) | (0.1168) | (0.1378) | (0.1817) | (0.1505) | (0.1450) | (0.2505) | (0.1783) |
| Children in HH | 0.1545 | 0.0353 | -0.0989 | -0.1540 | 0.2173 | 0.0986 | -0.2234 | 0.2146 | -0.0663 |
| | (0.1197) | (0.1108) | (0.1156) | (0.1404) | (0.1781) | (0.1487) | (0.1421) | (0.2353) | (0.1687) |
| Age (Decades) | -0.2100*** | -0.0142 | 0.0335 | -0.1625*** | -0.0998 | -0.0284 | -0.0332 | -0.3256^{***} | 0.1005^{*} |
| | (0.0404) | (0.0362) | (0.0402) | (0.0462) | (0.0609) | (0.0482) | (0.0473) | (0.0976) | (0.0539) |
| College | -0.0078 | -0.1608 | -0.2655** | 0.0753 | 0.2799 | 0.1874 | 0.0563 | 0.3214 | -0.1017 |
| | (0.1286) | (0.1181) | (0.1237) | (0.1470) | (0.1887) | (0.1584) | (0.1524) | (0.2427) | (0.1861) |
| HH Income (10k) | -0.0054 | 0.0011 | -0.0192* | -0.0228* | 0.0187 | 6.3×10^{-5} | -0.0018 | -0.0370 | 0.0101 |
| | (0.0115) | (0.0104) | (0.0111) | (0.0136) | (0.0157) | (0.0141) | (0.0142) | (0.0263) | (0.0178) |
| Rural | 0.4936*** | 0.1540 | -0.2102 | 0.1996 | -0.1459 | 0.0325 | -0.0731 | -0.5656 | 0.2309 |
| a , , | (0.1786) | (0.1632) | (0.1757) | (0.2164) | (0.2524) | (0.2267) | (0.2229) | (0.3539) | (0.2690) |
| Suburban | 0.0081 | 0.0997 | -0.2578* | 0.3706** | -0.1169 | 0.1036 | 0.1204 | -0.5205* | 0.1786 |
| | (0.1508) | (0.1343) | (0.1440) | (0.1826) | (0.1893) | (0.1848) | (0.1841) | (0.2681) | (0.2153) |
| White Evangelical | 0.3087** | -0.0887 | -0.2282* | 0.0114 | 0.6919*** | 0.1734 | -0.2617 | 0.6300** | -0.4650** |
| Democrat | (0.1287) | (0.1198) | (0.1282) | (0.1490) | (0.2008) | (0.1625) | (0.1623) | (0.2790) | (0.2095) |
| Democrat | (0.1700) | -0.0527 | 0.0268 | -0.2203 | 0.0115 | -0.2277 | 0.2811 | -0.0531 | -1.013 |
| Danach liann | (0.1700) | (0.1556) | (0.1004) | (0.2000) | (0.2267) | (0.2152) | (0.2090) | (0.3000) | (0.2207) |
| Republican | (0.2934) | (0.0234) | -0.0525 | (0.4292) | -0.2342 | (0.2052) | 0.0908 | -0.4555 | -0.0013 |
| Othen Bents | (0.1392) | (0.1445) | (0.1340) | (0.1794) 0.4044* | (0.2444) | (0.1973) | (0.2010) | (0.2970) 0.1072 | (0.2109) 0.4549* |
| Other Farty | -0.3000 | -0.0462 | (0.3103) | (0.4944) | (0.2787) | 0.0300 | (0.2559) | -0.1075 | (0.2682) |
| Ideological Identity | (0.2001) | 0.0061 | 0.2043) | (0.2074) 0.1937*** | 0.0008 | 0.0674 | 0.1154** | 0.0861 | 0.1084* |
| Ideological Identity | (0.0387) | -0.0001 | (0.2043) | (0.0451) | (0.0562) | (0.0496) | (0.0480) | (0.0753) | -0.1084 |
| Region: Rockies | 0.3155 | 0.1459 | 0.3415 | 1 108*** | 0.5827 | 0.0490) | 0.3603 | 0.1076 | 0.0319 |
| Region. Rockies | (0.2916) | (0.2743) | (0.2982) | (0.3714) | (0.5117) | (0.4293) | (0.3747) | (0.7781) | (0.4068) |
| Region: Southwest | -0 4277 | 4.28×10^{-5} | 0.3089 | -0.3014 | -0.0541 | 0.3983 | 0.0417 | -0.3127 | 0.2605 |
| riegion. Southwest | (0.2733) | (0.2414) | (0.2553) | (0.2927) | (0.3554) | (0.3523) | (0.3226) | (0.6412) | (0.3519) |
| Region: Great Plains | -0.0964 | -0.3823 | 0.4861^* | -0.2683 | -0.5197 | 0.4211 | -0.3516 | 0.2430 | -0.2190 |
| regioni creat i lanie | (0.2922) | (0.2652) | (0.2880) | (0.3215) | (0.4591) | (0.3827) | (0.3622) | (0.6597) | (0.4003) |
| Region: Midwest | -0.0594 | -0.1353 | 0.1371 | -0.3769 | -0.1076 | 0.1562 | -0.0685 | 0.2731 | -0.1347 |
| | (0.2477) | (0.2197) | (0.2279) | (0.2682) | (0.3330) | (0.3406) | (0.2960) | (0.5405) | (0.3332) |
| Region: South | -0.0892 | -0.4891** | 0.3112 | -0.6602*** | -0.1674 | 0.2051 | -0.2067 | 0.5965 | -0.3131 |
| 0 | (0.2332) | (0.2080) | (0.2141) | (0.2550) | (0.3137) | (0.3258) | (0.2855) | (0.5092) | (0.3124) |
| Region: Mid-Atlantic | -0.0738 | -0.1591 | 0.2181 | -0.4103 | -0.1774 | 0.3126 | -0.1687 | -0.0481 | -0.5694 |
| 0 | (0.2846) | (0.2523) | (0.2625) | (0.3126) | (0.3687) | (0.3706) | (0.3457) | (0.6192) | (0.4174) |
| Region: New England | -0.3338 | -0.0601 | 0.3822 | -0.8598** | 0.0521 | 0.3755 | -0.2547 | 0.2420 | -0.3517 |
| | (0.3263) | (0.2855) | (0.3207) | (0.3795) | (0.4387) | (0.4108) | (0.3990) | (0.7311) | (0.4349) |
| Hardship Index | 0.5061 | 0.4168 | 0.4779 | 0.5863 | 0.7705^{*} | 1.112*** | 0.3885 | 1.339^{**} | -0.0924 |
| | (0.3389) | (0.3138) | (0.3196) | (0.3936) | (0.4438) | (0.4023) | (0.4315) | (0.6362) | (0.5702) |
| COVID in HH | -0.1215 | -0.1121 | -0.3377*** | -0.0173 | 0.2313 | 0.1455 | 0.0736 | -0.2911 | -0.2052 |
| | (0.1179) | (0.1086) | (0.1140) | (0.1378) | (0.1785) | (0.1494) | (0.1465) | (0.2362) | (0.1733) |
| Fit statistice | | | | * | | | | | |
| Observations | 1.810 | 1.810 | 1.810 | 1.810 | 1.810 | 1.810 | 1.810 | 1.810 | 1.810 |
| Squared Correlation | 0 10034 | 0.05635 | 0.00202 | 0.06008 | 0.07377 | 0.03514 | 0.03480 | 0.03867 | 0.04399 |
| Pseudo R ² | 0.09653 | 0.04378 | 0.07267 | 0.06545 | 0.09102 | 0.03687 | 0.03685 | 0.09287 | 0.06528 |
| BIC | 2.238.6 | 2.536.8 | 2.319.0 | 1.806.3 | 1.342.4 | 1.678.0 | 1.723.9 | 878.24 | 1.397.9 |
| | _, | _, | _, | -, | -, | -, | -,. 2010 | | -, |

Table C12: Reasons for Pandemic Gun Purchases (New Owners vs. Pre-Existing Owners)

| Dependent Variables: Model: | Hobbyist Reasons OLS | Threat Reasons OLS | Other OLS |
|---------------------------------------|-------------------------|-----------------------|----------------|
| Variables | | | |
| (Intercept) | 0.7102^{***} | 0.7212^{***} | 0.1748^{***} |
| · · · · · · · · · · · · · · · · · · · | (0.0737) | (0.0660) | (0.0519) |
| New Owner | -0.2709*** | 0.0985*** | -0.0170 |
| ITON OWINI | (0.0283) | (0.0216) | (0.0188) |
| Plack | -0.2020*** | 0.0491 | 0.0238 |
| Diack | -0.2020 | (0.0221) | (0.0201) |
| A | (0.0410) | (0.0321) | 0.0291) |
| Asian | -0.0114 | -0.0102 | 0.0209 |
| Hispanic | (0.0699) | (0.0570) | (0.0452) |
| | -0.1407*** | -0.0588 | 0.0260 |
| | (0.0517) | (0.0460) | (0.0375) |
| Other Race | 0.0705 | 0.0303 | 0.0287 |
| | (0.0644) | (0.0511) | (0.0515) |
| Female | -0.1201*** | -0.0272 | 0.0377^{**} |
| | (0.0254) | (0.0209) | (0.0164) |
| Children in HH | 0.0306 | -0.0086 | -0.0054 |
| | (0.0250) | (0.0207) | (0.0170) |
| Age (Decades) | -0.0213*** | -0.0123* | 0.0092* |
| nge (Decades) | (0.00210 | (0.0070) | (0.0052) |
| Collogo | 0.0060 | 0.0070) | 0.0005 |
| College | -0.0002 | -0.0084 | -0.0095 |
| HH Income (10k) | (0.0267) | (0.0221) | (0.0183) |
| | 0.0023 | -0.0035 | 0.0011 |
| | (0.0024) | (0.0021) | (0.0017) |
| Rural | 0.0559 | -0.0592^{*} | 0.0227 |
| | (0.0362) | (0.0306) | (0.0247) |
| Suburban | 0.0091 | -0.0548^{**} | 0.0168 |
| | (0.0296) | (0.0246) | (0.0189) |
| White Evangelical | 0.0290 | 0.0093 | -0.0356** |
| 0 | (0.0272) | (0.0236) | (0.0165) |
| Democrat | 0.0551 | 0.0415 | -0.0993*** |
| Domoorat | (0.0347) | (0.0294) | (0.0232) |
| Ropublican | 0.0480 | 0.0204) | 0.0671*** |
| Republican | (0.0921) | (0.0208 | -0.0071 |
| Other Bester | (0.0551) | (0.0275) | (0.0232) |
| Other Party | -0.0003 | 0.0477 | 0.0830 |
| T 1 1 1 T 1 1 1 | (0.0522) | (0.0412) | (0.0462) |
| Ideological Identity | 0.0012 | 0.0310*** | -0.0108** |
| | (0.0080) | (0.0068) | (0.0050) |
| Region: Rockies | 0.0293 | -0.0036 | 0.0002 |
| | (0.0638) | (0.0570) | (0.0451) |
| Region: Southwest | -0.0056 | 0.0156 | 0.0273 |
| | (0.0548) | (0.0484) | (0.0398) |
| Region: Great Plains | -0.0537 | 0.0531 | -0.0259 |
| - | (0.0603) | (0.0518) | (0.0413) |
| Region: Midwest | -0.0110 | 0.0122 | -0.0170 |
| negioni interiori | (0.0502) | (0.0443) | (0.0344) |
| Region: South | -0.0573 | 0.0411 | -0.0346 |
| reson. south | (0.0472) | (0.0416) | (0.0324) |
| Porion, Mid Atlantia | 0.0472) | 0.0941 | 0.0524) |
| negion. mid-Atlantic | 0.0012 | 0.0241 | -0.0007 |
| D · N D · · | (0.0572) | (0.0502) | (0.0359) |
| Region: New England | -0.0118 | 0.0323 | -0.0371 |
| Hardship Index | (0.0659) | (0.0588) | (0.0445) |
| | 0.1539^{**} | 0.1591^{***} | -0.0066 |
| | (0.0689) | (0.0520) | (0.0496) |
| COVID in HH | -0.0264 | -0.0194 | -0.0179 |
| | (0.0247) | (0.0205) | (0.0169) |
| E11 .4.47.47. | · / | · / | . / |
| Fit statistics | | | |
| Observations | 1,810 | 1,810 | 1,810 |
| Squared Correlation | 0.11066 | 0.03958 | 0.04774 |
| Pseudo R ² | 0.08121 | 0.03825 | 0.09047 |
| BIC | 2,604.0 | 2,040.4 | 1,092.7 |

Table C13: Aggregated Reasons for Pandemic Gun Purchases (New Owners vs. Pre-Existing Owners)

| Dependent Variables: Model: | Hobbyist Reasons Logit | Threat Reasons Logit | Other Logit |
|--------------------------------|---------------------------|-------------------------|-------------------|
| Variables | | | - |
| (Intercept) | 0.9300*** | 0.9782^{**} | -1.640*** |
| (intercept) | (0.3370) | (0.4013) | (0.5017) |
| New Owner | -1 206*** | 0 7087*** | -0 1528 |
| ivew Owner | (0.1364) | (0.1777) | (0.2000) |
| Black | -0.9388*** | 0.3730 | 0.2756 |
| Such | (0.2017) | (0.2490) | (0.2801) |
| Asian | -0.0503 | -0.1138 | 0.2031) |
| | -0.0000 | (0.2257) | (0.2020) |
| Hispanic | 0.6215*** | 0.2504 | 0.9545 |
| | -0.0313 | -0.3594 | (0.2040) |
| Other Race | (0.2341) | (0.2002) | 0.0122) |
| | 0.3107 | 0.2108 | 0.2299 |
| Female | (0.2884) | (0.3758) | (0.3866) |
| | -0.5401*** | -0.1840 | 0.4251^{**} |
| | (0.1144) | (0.1346) | (0.1783) |
| Children in HH | 0.1333 | -0.0496 | -0.0663 |
| | (0.1122) | (0.1328) | (0.1687) |
| Age (Decades) | -0.0975^{***} | -0.0800* | 0.1005^{*} |
| | (0.0363) | (0.0442) | (0.0539) |
| College | -0.0263 | -0.0537 | -0.1017 |
| | (0.1192) | (0.1384) | (0.1861) |
| HH Income (10k) | 0.0102 | -0.0206 | 0.0101 |
| | (0.0109) | (0.0127) | (0.0178) |
| Bural | 0.2509 | -0.3865** | 0.2309 |
| rearon | (0.1666) | (0.1971) | (0.2690) |
| Suburban | 0.0385 | -0.3647** | 0.1786 |
| Suburban | (0.1372) | (0.1652) | (0.2152) |
| White Examplies | (0.1373) | (0.1055) | 0.4650* |
| white Evangencal | (0.1021 | 0.0744 | -0.4050 |
| D | (0.1222) | (0.1470) | (0.2095) |
| Democrat | 0.2515 | 0.2571 | -1.013*** |
| Republican | (0.1580) | (0.1789) | (0.2287) |
| | 0.2175 | 0.1316 | -0.6013** |
| Other Party | (0.1477) | (0.1670) | (0.2109) |
| | -0.2942 | 0.3179 | 0.4542^{*} |
| | (0.2309) | (0.2802) | (0.2682) |
| Ideological Identity | 0.0063 | 0.1928^{***} | -0.1084° |
| | (0.0372) | (0.0435) | (0.0554) |
| Region: Rockies | 0.1231 | -0.0145 | 0.0319 |
| | (0.2922) | (0.3193) | (0.4068) |
| Region: Southwest | -0.0274 | 0.0820 | 0.2605 |
| | (0.2498) | (0.2870) | (0.3519) |
| Region: Great Plains | -0.2478 | 0.3245 | -0.2190 |
| region. Great I lams | (0.2709) | (0.3162) | (0.4003) |
| Region: Midwest | -0.0518 | 0.0607 | -0 1347 |
| region. midwest | (0.2297) | (0.2594) | (0.3332) |
| Domion, Couth | 0.2231) | 0.2482 | 0.2121 |
| Region. South | -0.2040 (0.2147) | (0.2465) | -0.3131 |
| | (0.2147) | (0.2457) | 0.5124 |
| Region: Mid-Atlantic | 0.0081 | 0.1416 | -0.5694 |
| | (0.2629) | (0.2976) | (0.4174 |
| Region: New England | -0.0514 | 0.2186 | -0.3517 |
| | (0.2945) | (0.3421) | (0.4349) |
| Hardship Index | 0.7205^{**} | 1.052^{***} | -0.0924 |
| | (0.3295) | (0.3746) | (0.5702) |
| COVID in HH | -0.1188 | -0.1237 | -0.2052 |
| | (0.1107) | (0.1288) | (0.1733) |
| Fit statistics | | | |
| Chammation - | 1 010 | 1 010 | 1 010 |
| Observations | 1,810 | 1,810 | 1,810 |
| Squared Correlation | 0.11085 | 0.04017 | 0.04322 |
| rseudo R [*] | 0.08442 | 0.04031 | 0.06528 |
| TO T (7) | 0 105 0 | O OOF F | 1 207 0 |

Table C14: Aggregated Reasons for Pandemic Gun Purchases (New Owners vs. Pre-Existing Owners)