



School Accountability and Teacher Mobility

Li Feng

Finance and Economics
Texas State University

David Figlio

Faculty Fellow, Institute for Policy Research
Education and Social Policy, School of Education and Social Policy
Northwestern University
Research Associate, National Bureau of Economic Research

Tim Sass

Economics
Florida State University

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Abstract

This paper presents the first causal evidence on the effects of school accountability systems on teacher labor markets. The researchers exploit a 2002 change in Florida's school accountability system that exogenously shocked some schools to higher accountability grades and others to lower accountability grades, and measure whether teachers in shocked schools are more or less likely to move. Using microdata from the universe of Florida public school teachers, they find strong evidence that accountability shocks influence the teacher labor market; specifically, teachers are more likely to leave schools that have been downward shocked—especially to the bottom grade—and they are less likely to leave schools that have been upward shocked. They also find that accountability shocks influence the distribution of the measured quality of teachers (in terms of value added measures) who stay and leave their school, though the average differences are not large.

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I. Introduction

School accountability -- the process of evaluating schools on the basis of the performance of their students and holding them responsible for this performance -- is becoming increasingly prevalent around the world. Accountability systems are intended to solve the principal-agent problem in education, and the incentives that they provide to educators to improve student efficiency take several forms. These systems provide direct incentives, in the form of explicit rewards or sanctions associated with good or poor performance. In addition, many of the mechanisms for improvement involve social pressure, since a school's constituents have both educational and financial reasons to influence low-performing schools to improve. The financial reasons derive from the fact that school accountability ratings tend to be capitalized into housing values (Figlio and Lucas, 2004). In addition, school accountability affects a school's ability to raise voluntary contributions (Figlio and Kenny, 2009). This paper makes use of detailed data at the individual teacher level to gauge the degree to which teachers respond to these direct and indirect forms of accountability pressure by leaving the high-pressure school.

There is certainly reason to believe that educators respond to accountability pressure in the ways in which they carry out their jobs. Early evidence concerning the effects of these accountability systems on student performance indicates they tend to improve the outcomes of low-performing students (see, e.g., Chakrabarti, 2006; Chiang, 2009; Figlio and Rouse, 2006; Jacob, 2005; Rouse et al., 2007; West and Peterson, 2006).¹ However, Krieg (2008), Neal and Schanzenbach (2008) and Reback (2008) argue that the benefits of accountability pressures are concentrated in the more marginal students rather than the students whose performance would be

¹ Recent nationwide studies by Carnoy and Loeb (2002) and Hanushek and Raymond (2005) also find significant improvement in student outcomes as a result of standards-based accountability, whereas the results from some specific state systems have been less positive (see, e.g, Koretz and Barron, 1998; Clark, 2003; and Haney, 2000, 2002).

far above or far below the performance thresholds set for the purposes of school accountability. That said, the weight of the evidence suggests that accountability systems have led schools to become more productive, at least along some measurable dimensions. Rouse et al. (2007) document a number of the ways in which accountability pressure has changed school instructional policies and practices in Florida's low-performing schools, and relate these instructional policy and practice changes to increased student performance.

In addition to actively changing policies and practices to improve student outcomes, schools have also responded to accountability pressures by engaging in apparently strategic behavior with questionable educational benefit. For instance, some schools have responded by differentially reclassifying low-achieving students as learning disabled so that their scores will not count against the school in accountability systems (see, e.g., Cullen and Reback, 2007; Figlio and Getzler, 2007; Jacob, 2005).² Figlio and Winicki (2005) suggest that Virginia schools facing accountability pressures altered their school nutrition programs on testing days to increase the likelihood that students will do well on the exams, and Figlio (2006) indicates that schools differentially suspend students at different points in the testing cycle in an apparent attempt to alter the composition of the testing pool. Jacob and Levitt (2003) find that teachers are more likely to cheat when faced with more accountability pressure. And the distributional effects documented by Neal and Schanzenbach (2008) and others are also evidence of strategic behavior on the part of schools. In sum, school accountability systems cause schools to behave differently, and school personnel almost certainly are very responsive to increased accountability pressure.

² Chakrabarti (2006), however, does not find that schools respond in this way.

Of course, the individuals implementing the changes in instructional policies and practices are teachers, and school accountability therefore has the potential to influence the desirability of certain teaching jobs. Likewise, accountability may influence the willingness of schools to retain certain teachers. From a theoretical perspective, the effects of accountability pressures on the teacher labor market are ambiguous. On the demand side, in order to avoid sanctions and/or the stigma associated with being designated as a “failing” school, schools could increase their efforts to identify low performing teachers and remove them from their classrooms. In this case, it is difficult to call these personnel changes a “job choice,” at least from the perspective of the teacher. On the supply side, accountability pressure and associated changes in school policies could lower the net benefit of teaching in a school by reducing teacher discretion over curriculum or teaching methods. Likewise, the potential stigma from teaching in a “failing” school could lead some teachers to seek employment at other schools. On the other hand, the resources that often accompany sanctions (e.g. reading coaches, training for teachers, etc.) could reduce the non-monetary costs associated with working in low-performing schools and actually increase teacher retention.

Similar patterns could be possible in the case of schools receiving high accountability marks. Schools that perform well under accountability systems face the pressure to maintain their high scores or to improve upon them. In Florida, for instance, schools receive extra bonus money for maintaining a top accountability score or for improving from one year to the next. Given that the housing market capitalization effects of school accountability are strongest in more affluent school areas (Figlio and Lucas, 2004), coupled with the fact that measurement error in test scores introduces a significant degree of randomness into the school accountability ratings (Kane and Staiger, 2004), it is reasonable to believe that relatively high-performing

schools may face as much or more real accountability pressure as do low-performing schools. Goldhaber and Hannaway (2004), in a case study of schools in Florida, find that teacher and administrator anxiety levels due to school accountability were highest in the high-performing schools, where school personnel felt the most pressure to maintain their accountability marks. Hence, while teaching in a highly-rated school has its advantages, the extra accountability pressure may deter teachers as well. And, just as with the low-rated schools, highly-rated schools may engage in more teacher selection as a consequence of accountability. Therefore, the theoretical expectations of how accountability pressures influence teacher job choice/placement are ambiguous for both highly-rated and low-rated schools.

A number of recent papers have analyzed the determinants of teacher mobility and attrition (Boyd et al., 2005a, 2006, 2007; Clotfelter et al., forthcoming; Feng, 2009; Hanushek et al., 2004; Imazeki, 2004; Krieg, 2006; Podgursky et al., 2004; Scafidi et al. 2007). However, the literature relating accountability pressures to teachers' labor market decisions has been much spottier. Boyd et al. (2005) explore the responses of teachers to the introduction of mandated state testing in New York State. They find that teacher turnover in fourth grade, the critical accountability year in New York, decreased following the introduction of testing, and that entrants into fourth grade were more likely to be experienced than had previously been the case. Clotfelter et al. (2004) evaluate how North Carolina's accountability system has influenced the ability of schools serving low-performing students to attract and retain high-quality teachers. They find that the introduction of the accountability system has exacerbated teacher turnover in these schools, though it is less evident that accountability has led to lower qualifications of the teachers serving low-performing students. Both of these papers carefully describe the

accountability systems in their states, but because they evaluate accountability systems that affected all schools within a state, it is difficult to derive causal inference from their analyses.

In this paper, we exploit a major rule change in Florida's school accountability system that took place in the summer of 2002 to identify the effects of changing school accountability pressures on teacher job changes. Florida had graded every school in the state on a scale from "A" to "F" since the summer of 1999, based on proficiency rates in reading, writing and mathematics. In 2002, the state dramatically changed its grading system to both recalibrate the acceptable student proficiency levels for the purposes of school accountability and to introduce student-level changes as an important determinant of school grades. Using student-level micro-data to calculate the school grades that would have occurred absent this change, we demonstrate that over half of all schools in the state experienced an accountability "shock" due to this grading change, with some schools receiving a higher grade than they would have otherwise received and other schools receiving a lower grade than would have otherwise occurred. Furthermore, some schools were shocked downward to receive a grade of "F", which no school in the state had received in the prior year of grading. These grading shocks provide the vehicle for identification of accountability effects in this paper.

We apply these accountability shocks to data on the universe of public school teachers in Florida. We measure the effects of accountability pressures on teachers' decisions to stay at a given school, move to another school in the same district, move to another school district in the state or leave public school teaching. Since Florida has had statewide achievement testing in all grades 3-10 since 1999/2000 we are also able to compute "value-added" measures of teacher quality and determine whether teacher mobility engendered by accountability pressures tends to increase or decrease teacher quality. We find strong evidence that accountability shocks

influence the teacher labor market; specifically, teachers are more likely to leave schools that have been downward shocked -- especially to the bottom grade -- and they are less likely to leave schools that have been upward shocked. We also find that accountability shocks influence the distribution of the measured quality of teachers (in terms of value added measures) who stay and leave their school, though the average differences are not large. The results, therefore, suggest that school accountability can have quite consequential effects for professional educators.

II. The Florida School Accountability Program

Education reform, and specifically a system of school accountability with a series of rewards and sanctions for high-performing and low-performing schools, was the policy centerpiece of Jeb Bush's 1998 gubernatorial campaign in Florida; the resulting A+ Plan for Education was Bush's first legislative initiative upon entering office in 1999. The A+ Plan called for annual curriculum-based testing of all students in grades three through ten, and annual grading of all traditional public and charter schools based on aggregate test performance. As noted above, the Florida accountability system assigns letter grades ("A," "B," etc.) to each school based on students' achievement (measured in several ways). High-performing and improving schools receive rewards while low-performing schools receive sanctions as well as additional assistance, through Florida's Assistance Plus program. The most famous and publicized provision of the A+ Plan was the institution of school vouchers, called "Opportunity Scholarships," for students attending (or slated to attend) chronically failing schools -- those receiving a grade of "F" in two years out of four, including the most recent year. Opportunity Scholarships allowed students to attend a different public school, or an eligible private school.

School grading began in May 1999, immediately following passage into law of the A+ Plan. In each year from 1999 through 2001, a school would earn a grade of “F” if fewer than 60 percent of students scored at level 2 (out of 5) or above in reading, fewer than 60 percent of students scored at level 2 (out of 5) or above in mathematics, and fewer than 50 percent of students scored at level 3 (out of 6) or above on the Florida Writes! writing evaluation, known from 2001 onward as the FCAT Writing examination. A school could avoid the “F” label by meeting any one of these three standards in 1999; the same was true in 2000 and 2001 provided that at least 90 percent of the test-eligible students took the examination (or that the school could provide the state with a “reasonable explanation” for why fewer than 90 percent of students took the test.) All schools in the distribution were subject to accountability pressure, and according to Goldhaber and Hannaway (2004), schools throughout the distribution apparently felt pressure to perform in measurable ways.

Thus, between 1999 and the summer of 2001, schools were assessed primarily on the basis of aggregate test score *levels* (and also some additional non-test factors, such as attendance and suspension rates, for the higher grade levels) and only in the grades with existing statewide curriculum-based assessments,³ rather than on *progress* schools make toward higher levels of student achievement. Starting in summer 2002, however, school grades began to incorporate test score data from all grades from three through ten, and were also the first to evaluate schools not just on the level of student test performance but also on the year-to-year progress of individual students. In our analysis, we take advantage of the fact that during the 2001-02 school year teachers would not have been able to anticipate their school grade in summer 2002 because of the changes in the formula and because the changes were not decided until the last minute.

³ Students were tested in grade 4 in reading and writing, in grade 5 in mathematics, in grade 8 in reading, writing and math, and in 10 in reading, writing and math.

By the beginning of the 2001-02 school year several things were known about the school grades that were to be assigned in summer 2002. First, the school grades were to have been based on the test scores of all students in grades three through ten in reading and mathematics, and in the fourth, eighth and tenth grades in writing. Second, the standards for acceptable performance by students were to be raised from level 2 to level 3 in reading and mathematics. Third, some notion of a “value-added” system was to be established, though little was known about the specific nature of this system except that it would augment the levels-based grading system and would focus principally on the lowest-performing students. These elements would be combined to give each school a total number of “grade” points. The school’s grade would be determined by the number of points. However, the specifics of the formula that would put these components together to form the school grades were not announced until March 5, 2002, leaving teachers entering the 2001-02 school year with virtually no information with which to anticipate their school’s exact grade.

Table 1 shows the distribution of schools across the five performance grades for the first six rounds of school grading, for all graded schools in Florida. As is apparent from the variation across years in the number of schools that fall into each performance category, there are considerable grade changes that have taken place since the accountability system was adopted. Most notable is the fact that while 70 schools received an F grade in the first year (1998-99) only 4 did so the subsequent year and none did by the summer of 2001. At the same time, an increasing number of schools were receiving A’s and B’s. This is partly due to the fact that schools had learned their way around the system: A school had to fail to meet proficiency targets in all three subjects to earn an F grade so as long as students did well enough in at least one subject the school would escape the worst stigma. Hannaway and Goldhaber (2004) and

Chakrabarti (2006) find evidence that students in failing schools made the biggest gains in writing which is viewed as one of the easier subjects in which to improve quickly. When the rules of the game changed, so did the number of schools caught by surprise: For instance, 60 schools earned an "F" grade in the summer of 2002. The number of schools that received an A grade also increased, due in large measure to the shift to the "grade points" system of school grading, which allows schools that miss performance goals in one area to compensate with higher performance in another area. Finally, note that as schools have adapted to the new grading system, the number of failing schools has decreased.⁴

In this paper, we seek to exploit the degree to which schools and teachers were "surprised" by the change in school grading. Using an approach to identification introduced by Rouse et al. (2007), we measure the "accountability shock" to schools and teachers by comparing the grades that schools actually received in the summer of 2002 with the grade that they would have been predicted to receive in 2002 based on the "old" grading system (that in place in 2001). We have programmed both the old and new accountability formulas and, using the full set of student administrative test scores provided us by the Florida Department of Education, we have calculated both the actual school grade that schools received in 2002 with the grade that they would have received given their realized student test scores had the grading system remained unchanged. It is essential that we make this specific comparison, rather than simply comparing year-to-year changes in school grades, because year-to-year changes in school grades could

⁴ Note that in Table 1 there are 68 elementary schools that received a grade of "N" in 2002. These were new schools in that year. As such, they were not given a formal grade although the state did calculate their accountability points. Rouse et al. (2007) experimented with imputing what their grades would have been, and found that there would have been an additional 9 "F" graded schools and 10 additional "D" graded schools, for instance, had the state graded these schools in 2002. Rouse et al. (2007) found that their results regarding test scores were completely insensitive to the treatment of these schools. For the purposes of the present analysis, we exclude them from the analysis of teacher job changes.

reflect not just accountability shocks, but also changes in student demographics, changes in school policies and practices, changes in school staffing, and other changes in school quality. Given that understanding school staffing is the point of this paper, it is clearly inappropriate to compare grade changes per se.

Table 2 compares realized school grades to predicted school grades (based on the old grading system but the new student test scores) for the set of schools in the state of Florida.⁵ We demonstrate that 51 percent of schools experienced a change in their school grade based on the changing parameters of the grading system itself. Most of these schools (42 percent) experienced an upward shock in their school grades, while 9 percent of all schools experienced a downward shock in their school grades, receiving a lower grade than they would have expected had the grading system remained unchanged. Twenty percent of schools that might have expected to receive a “D” under the old system received an “F” under the new one, while 38 percent of these schools received a grade of “C” or better. Meanwhile, 59 percent of schools that might have expected to receive a “B” under the old system received an “A” under the new one, while 9 percent of these schools received a grade of “C” or worse. It is clear that the grading system change led to major changes in the accountability environment, and provides fertile ground for identification.

⁵ The number of observations in Table 2 does not exactly match that in Table 1 because we rely on administrative data on students provided by the Florida Department of Education to simulate each school’s grade in 2002. This administrative dataset does not include some students in charter schools, “alternative” schools, and, of course, schools that do not have any students in the accountability grades.

III. Data and Empirical Approach

We are interested in modeling the effects of school accountability shocks on teacher mobility. The most natural way to estimate this relationship is to consider year-to-year changes in teacher employment at a school. In this paper, we estimated log-odds ratios generated from logit models in which the dependent variable is the likelihood that a teacher in year t leaves his or her school before year $t+1$, and our key independent variables are indicators for whether the school was upward or downward shocked in the 2002 school grading regime -- a change in accountability pressure that is exogenous to the school and its teachers. As we observe teacher mobility decisions in the years before and after the school grading change, we can estimate difference-in-difference models in which we condition on school and time effects. Our models are estimated at the individual teacher level, but since our treatment is a school-level treatment, all standard errors are clustered at the school level. We report the results of models both including and excluding school and student body characteristics.

The primary source of our data is the Florida Department of Education's K-20 Education Data Warehouse (FL-EDW), an integrated longitudinal database covering all Florida public school students and school employees from pre-school through college. Like statewide administrative databases in North Carolina and Texas, the FL-EDW contains a rich set of information on both individual students and their teachers which is linked through time. Unlike other statewide databases, however, the FL-EDW links both students and teachers to specific classrooms at all grade levels.

Statewide data, as opposed to data from an individual school district, are particularly useful for studying teacher labor markets since we can follow teachers who move from one district to another within Florida. We cannot, however, track teachers who move to another

state. Due to population growth and a constitutionally mandated maximum class size, Florida is a net importer of teachers. Thus, unlike many Northern states where the school-age population is shrinking, there is relatively little outflow of teachers from Florida during the period under study.⁶ Using national data from the Schools and Staffing Survey and associated Teacher Follow-Up Survey (SASS/TFS) which track teachers across state lines, Feng (2010) finds that there are relatively fewer teachers moving into or out of state of Florida compared to other Southern states, such as North Carolina and Georgia.

The FL-EDW contains teachers' data from the 1995/96 school year to the 2004/05 school year, though our primary analyses are based on data through 2002/03, the first year following the school accountability shock. Teachers' school affiliation and status can be identified for the universe of all classroom teachers. In some of our specifications we condition on student test performance (or measure teachers' value added based on student performance); student test score records for all grades 3-10 are only available from the 1999/2000 school year forward. We also, in some specifications, we control for average test scores, disciplinary incidents and socio-economic status and demographics of the students in a teacher's school and district in an attempt to control for other factors that might affect teacher mobility.

IV. Accountability Shocks and Teacher Mobility

We begin by investigating the number of relevant teachers who faced different accountability conditions during the accountability shock of summer 2002. Table 3 presents a descriptive summary of teacher job change before and after the accountability shock, broken

⁶ With the recent sub-prime mortgage crisis after 2008, one might expect more of a teacher exodus out of state of Florida than has occurred previously.

down by the type of shock a school received.⁷ Independent of the type of accountability shock, teacher mobility was increasing over the time period. This general increase in teacher job change could be due to the general expansion in the number of teachers employed; relatively inexperienced teachers tend to be more mobile than veteran teachers. Prior to the accountability shock, downward shocked schools have the highest teacher departure rate (18.2 percent), with slightly lower rates in no-shock schools (16.6 percent) and upward shocked schools (15.9 percent). There is little difference in the inter-temporal change in teacher job change between upward-shock and no-shock schools. The fraction of teachers leaving non-shock schools was 0.7 percentage points higher in the post-accountability-shock period and for upward-shock schools it was 1.0 percentage points higher. However, there is a relatively large and statistically significant difference in the change in teacher departures between no-shock and downward-shock schools. Teacher job change increased by 2.8 percentage points in schools that received lower grades under the new accountability regime, whereas schools that did not experience a change in their grade as a result of the change in the accountability system had only one percentage point increase in teacher mobility. This 1.8 percentage point difference in the rate of teacher departure is statistically significant at better than a 99.9 percent confidence level.

The simple descriptive evidence suggests that schools experiencing downward shocks see more of their teachers depart after the implementation of the accountability formula change than do schools that did not experience an accountability shock. In contrast, there is no significant difference in teacher departure from schools that receive higher grades under the new

⁷ Since 2002 school grades were announced in mid-June 2002, too late for most teachers to make a job change before the start of school in August, we define the pre-shock period as job changes that occurred between the three school-year pairs, 1999/00-2000/01, 2000/01-2001/02 and 2001/02-2002/03. The post-shock period is the transition between school years 2002/03 and 2003/04.

accountability system relative to schools that experience no change in their grade as a result of the change in the accountability formula. It is possible that any improvement in teachers' perceptions of their work environment engendered by an unexpected increase in the school grade is offset by increased pressure to maintain a higher grade. Of course these descriptive results do not directly account for other time-varying factors that may influence teacher mobility and which could be correlated with the type of shock a school receives. Further, the descriptive comparisons are based on teacher job changes occurring between the pair of school years following the accountability shock (ie. differences between the 2002/03 and 2003/04 school years). If teacher job changes occur over a longer time frame, our results may underestimate the behavioral response to an accountability shock.

To account for other time varying factors that may affect teacher mobility, we first estimate logit models of the likelihood of leaving the current school, conditional on the characteristics of the teacher's current school. Results are presented in Table 4. For ease of interpretation, the table reports the parameters in terms of the odds ratio, relative to the reference group of teachers (those whose schools did not face an accountability shock.) An odds ratio of one indicates that a teacher whose school is shocked has equal odds of leaving the school as a teacher whose school is not shocked. An odds ratio of less than one implies that the odds of leaving for a teacher whose school is shocked is lower than for the reference group of teachers. Our key variable of interest is the accountability shock faced by the cohort of teachers. Because we observe multiple cohorts of teachers from schools, before versus after the accountability shock, we report differences-in-differences estimates in which we compare, say, teachers' decisions to move before versus after the policy shock from schools experiencing a downward

shock to those from schools that are un-shocked. All standard errors are adjusted for clustering at the school year level.

The first column of Table 4 presents the key parameter estimates of this model,⁸ in a specification that includes no covariates except for year indicators, a post-shock indicator, and shock-type fixed effects for upward and downward shocked schools. Similar to the descriptive statistics reported above, we find that teachers who are teaching in schools with an unexpected increase in accountability pressure (downward accountability shock) are about 11 percent more likely to leave their schools, and those with an unexpected decrease in accountability pressure (upward accountability shock) are 2.3 percent less likely to leave their schools, than are teachers facing no accountability shock. These estimated effects are statistically distinct from one another and the downward-shocked estimate is statistically different from the no-shock status at a 90 percent confidence level. Therefore, this simplest specification provides some initial evidence that teachers may respond to the information conveyed by accountability shocks.

The remaining columns of Table 4 include increasing numbers of control variables. The second column controls for time-varying district-level demographics and FCAT math score level variables, and the third column controls for both district-level and school-level time-varying demographic variables. These include disciplinary incidents, percent of minority students (Black, Hispanic) and percent of students receiving subsidized lunches, in addition to time effects and shock-level fixed effects.⁹ The fourth column further controls for school-level FCAT math score levels, and the fifth column controls for both school-level FCAT math score levels as well as math score gains. The point estimates are essentially constant across specifications, with the

⁸ For the sake of presentation parsimony we report only estimates of the key parameters of interest, but all coefficient estimates are available on request from the authors.

⁹ Because we are conceiving of the accountability shock as a school-level treatment, we do not condition on teacher characteristics.

precision of the shock effect estimates generally increasing with the number of controls in the model. Given the stability of our findings across model specifications, we focus on the fully specified model -- the model that controls for school and district student characteristics and school-level test score levels and gains (model 5) in the tables that follow. The descriptive results and parsimonious model 1 results are both supported by our preferred specification, model 5. Teachers teaching in downward shocked schools are 11 percent more likely to leave their schools compared to teachers in no-shock schools post reform. Differences between upward shocked schools and no-shock schools are not statistically significant throughout various specifications, suggesting teachers in these schools are just as likely to leave their schools as their counterparts in no-shock schools.

In the last two columns of Table 4 we report two conditional logit specifications. The first conditional logit model is similar to logit model 5 with the exception that shock-level indicators are now replaced by school fixed effects to better control for time-invariant school factors that may affect teacher mobility. Schools and teachers may react differentially to upward and downward accountability shocks that lead to different school grades. Figlio et al. (2007), for instance, find that test scores for "F" graded schools change more than test scores for other downward-shocked schools following the accountability policy change, and test scores decrease in "A" graded schools relative to those in other upward-shocked schools. We are interested in finding out whether teachers respond differently to being accountability-shocked to different portions of the grading distribution as well. The final column of Table 4, therefore, presents an augmented version of our most preferred specification broken out to reveal differences in teachers' responses by grade status.

Estimates in the final column of Table 4 show that teachers do indeed respond differently when shocked to different portions of the school grading distribution. Teachers in schools whose school grade unexpectedly falls to a "B", "C" or "D" are no more likely to leave their school than are teachers in schools that did not experience a change in grade due to the new accountability system. However, teachers in schools that experienced an unexpected drop into the failing category are 42 percent more likely to leave their school than are teachers in schools that did not experience an accountability shock. Once the difference between downward shock to "F" and downward shocks to higher grades are taken into account, teachers in schools that were upward shocked are found to be 6.5 percent less likely to leave their schools than teachers in no-shocked schools. However, the impact on teacher job change from being upwardly shocked to a school grade of "A" is not statistically different from that resulting from being upwardly shocked to a "B" or "C" (there were no "F" schools in the pre-shock year).¹⁰ Hence, it appears that teachers who receive unexpected good news about the rating of their school are moderately more likely to stay in their current school whereas teachers are much more likely to depart their school if it is unexpectedly hit with a failing grade.

Prior research on general teacher mobility distinguishes between moves between schools within a district, moves between districts and exit from public school teaching and finds that different factors can affect the type of move (c.f. Hanushek et al., 2004; Boyd, et al., 2005a; Scafidi, et al., 2007; Feng, 2009). These studies show that working conditions have a large impact on teacher mobility within a district, but inter-district moves and exits are more likely to

¹⁰ The coefficient of the differential effect of upward shock to "A" versus upward shock in general is actually positive, though statistically insignificant. While the sign of this differential relationship may seem surprising, it is consistent with the qualitative findings of Goldhaber and Hannaway (2004) that suggest that educators in "A" schools felt more under pressure to maintain their high grades than did educators in lower-rated schools.

be affected by factors other than working conditions, such as salary differentials and geographic preferences. The accountability pressure under study here can be considered as a characteristic of teachers' working conditions and thus we would expect there be differential effects of accountability pressure on the type of move. In particular, we would expect accountability pressure to have a greater effect on within district moves and exits than on cross-district job changes.

In Table 5 we present a set of conditional logit models by three possible transition destinations: intra-district move, inter-district move, and exit Florida public school teaching. The results are consistent with our general findings on the effect of accountability on teacher job change as well as the evidence in the general teacher mobility literature on the differential effects of working conditions on alternative types of teacher transitions. Schools shocked downward to "F" grades tend to lose teachers to other schools in the same district. Specifically teachers in these schools are nearly 67 percent more likely to move to a new school within the same district than are teachers in schools that experienced no accountability shock. Not surprisingly, movement of teachers between Florida's countywide school districts is not significantly affected by downward accountability shocks, including receipt of an "F" grade. Most likely such relatively long-distance moves are governed primarily by external factors such as spousal job changes or other significant family events, rather than a desire to change one's work environment. Being shocked down to a failing grade does increase the likelihood that a teacher will exit the Florida public school system entirely. Although the exit effect is smaller, the evidence suggests that teachers are both more likely to switch schools and to give up on public school teaching when their school is designated as "failing."

In contrast to downward accountability shocks, receipt of a higher school grade or upward movement to an “A” grade does not have a statistically significant effect on either within-district teacher mobility or on the probability of remaining as a public school teacher, though the estimated effects on leaving teaching in the school district (or entirely) are on the margin of statistical significance at conventional levels. Thus it appears that movements to the left tail of the school grade distribution have the greatest consequences for both the total supply of teachers and the distribution of teachers across schools.

VI. Teacher Quality and Differential Mobility

As numerous authors (e.g., Rivkin, Hanushek and Kain, 2005; Aaronson, Barrow and Sander, 2007) have demonstrated, there exists dramatic within-school variation in the level of teacher quality. The degree to which teacher mobility engendered by accountability pressures affects the level and distribution of teacher quality across schools depends critically on which teachers within a school stay or go. If relatively high-quality teachers depart schools facing accountability pressure this could mitigate any direct benefits to student learning brought about by the accountability system. In contrast, if increased pressure leads to (either voluntary or involuntary) exit of the least capable teachers this could reinforce the direct positive effects of accountability pressure.¹¹

While teacher quality is multidimensional, for the purposes of this paper we simply define teacher quality in terms of a teacher’s individual contribution to her student’s test scores. This is possible in the state of Florida because of the ability to link individual teachers to the

¹¹ For analyses of the effects of teacher mobility on teacher quality in non-accountability contexts see Goldhaber, Gross and Player (2007) and Feng and Sass (2008).

students for whom they are responsible on a classroom basis. Specifically, we estimate a so-called “restricted value-added” model of student achievement of the following form:¹²

$$\Delta A_{it} = \beta_1 Z_{it} + \beta_2 P_{-ijmt} + \beta_3 T_{kt} + \beta_4 S_{mt} + \delta_{kt} + v_{it} \quad (8)$$

The gain in student achievement, ΔA_{it} is a function of student/family inputs (Z), classroom-level peer characteristics (P), a vector of teacher experience indicators (T),¹³ school-level inputs (S) and a year-specific teacher effect, δ . The subscripts denote students (i), classrooms (j), teachers (k), schools (m) and time (t).¹⁴ The model is estimated for both math achievement and for reading achievement in all grades 3-10 over the period 1999/2000 – 2004/05.¹⁵ Only students with a single teacher in the relevant subject area are included in the analysis.

The estimated value of the year-specific teacher effect, δ_{kt} , is our measure of teacher quality. The teacher-by-year estimates are re-centered to have a mean value of zero in each school level (elementary, middle, high) within each year. The estimates represent the average achievement gain of a teacher’s students, for all classes taught in the relevant subject in a year, controlling for student, peer and school characteristics as well as for teacher experience. Student achievement is measured by year-to-year gains in the normalized-by-grade-and-year FCAT-NRT

¹² We also produced teacher quality estimates from two alternative models: an “unrestricted” value-added model whereby current achievement is regressed on lagged achievement and the same set of controls as in equation (8) and a “restricted” value-added model like (8) in which student fixed effects were used to control for student heterogeneity. We obtain very similar results for the relationship between accountability pressure and teacher quality using these alternative specifications. Results using these alternative value-added specifications are available from the authors.

¹³ We control for teacher experience, but not other teacher characteristics, because we are interested in identifying measured teacher quality for a given level of experience. We have also estimated models in which teacher experience levels are excluded, and the results are fundamentally similar to those reported herein.

¹⁴ For a derivation of the value-added model and its implicit assumptions, see Todd and Wolpin (2003).

¹⁵ Details of the estimation sample and estimation procedures are provided in Harris and Sass (2007).

score. Thus the teacher effects are calibrated in standard deviation units. Since there are no school-level fixed effects in the calculation of teacher effects, our year-specific teacher effect can be interpreted as the effectiveness of a given teacher relative to all other teachers teaching the same subject with the same level of experience at the same type of school.¹⁶ The teacher quality measure can only be constructed for teachers who are responsible for teaching courses in the subjects and grades covered by achievement testing, reading and math. In the following discussion we focus on the quality and mobility of math teachers. Results for reading teachers are very similar, and are available on request from the authors.

Table 6 reports the average teacher value added for stayers versus leavers in schools that were negatively shocked, those that were positively shocked, and those that were un-shocked when the accountability system changed. It appears that in the case of positively shocked and un-shocked schools, the average quality of teachers did not change following the change in school accountability. However, in the case of negatively shocked schools, the average teacher quality of leavers and stayers both improved after the shock. While the average quality of teachers who depart improved more than the average teacher quality of stayers the difference in the quality of leavers and stayers as a result of downward accountability shocks is not statistically significant.

In addition to comparing stayers and leavers within a given type of school, we also compare the average quality of leavers between accountability-shocked schools and unshocked schools in Table 7. The sample size is smaller than in Table 3 since some teachers are teaching in non-tested subjects or grade levels or can't be reliably matched to the performance of their students. As a consequence, results must be interpreted with some caution. For schools

¹⁶ We also analyzed teacher quality measured not conditioned on experience and obtained similar results.

experiencing either no shock or an upward shock, the average quality of leavers does not change significantly over time. In contrast, teachers who departed upwardly shocked schools were of significantly higher measured quality than those who left downwardly shocked schools before the accountability change. As a result, the difference between the quality of teachers who leave downwardly shocked schools and those who depart no-shock schools increased over time. Thus unexpected worsening of a school's grade lead to increases in both the number of teachers departing and the average quality of teachers who leave, relative to schools that did not experience a change in their school's grade.

In Table 8 we make a similar comparison between the quality of stayers in shocked and un-shocked schools. The average quality of stayers improved in downward shocked schools by about 0.05 standard deviations in student achievement while there was no significant change in the quality of teacher who stay in un-shocked schools. This is consistent with the findings of Rouse et al. (2007), suggesting that downward-shocked schools experience larger test score gains the year after the accountability shock. The apparent increase in stayer quality could be due to positive effects of accountability (either increased resources like reading coaches or increased pressure to perform) that improve teacher quality.

Table 9 demonstrates that the effects of downward accountability shocks on the quality of teachers who stay and leave essentially balance out. Unexpected declines in a school's grade lead to higher quality teachers departing but also lead to improvements in the quality of teachers who stay. The net result is that the difference between the quality of stayers and leavers in downward-shock schools and the difference between the quality of stayers and leavers in no-shock schools is unchanged. Thus it appears that increased accountability pressure has both

positive and negative effects on average teacher quality and that schools receiving unexpected decreases in school grades do not lose ground on the teacher quality margin.

The mean changes in teacher value added might mask changes in the distribution of movers and stayers. In order to understand the impact of accountability pressure on the distribution of teacher quality, we plot kernel density estimates of the distribution of teacher quality. Figure 1 illustrates the distribution of mathematics teacher quality before and after the accountability shock by move status for schools that received no change in accountability and for schools that experience increased pressure (downward shock) as a result of the change in the school grading system. In zero-shock schools there is little difference in the quality distributions of leavers and stayers, but the quality distribution of leavers shrinks after the accountability system change. Un-shocked schools experience fewer instances of teachers leaving at both ends of the quality spectrum. In schools with increased accountability pressure there are more pronounced differences in the quality distributions of teachers who stay and those that leave. Prior to the accountability system change, the quality distribution of movers is to the left of the stayer distribution and the right tail of the stayer quality distribution is longer. Following the accountability change, the right tail of the quality distribution of teachers staying in negatively-shocked schools is truncated to become nearly equal to the right tail of the leaver quality distribution. Further, the quality distribution of leavers shifts to the right while the distribution of stayers remains relatively constant over time.

VII. Conclusion

This paper provides new evidence of the effects of school accountability systems on teacher mobility. While prior papers on the subject analyzed the introduction of an

accountability system within a state, and therefore had no natural counterfactual, this study is the first to exploit policy variation within the same state to study the effects of accountability on teacher job changes.

Taking advantage of exogenous changes in accountability pressure for schools, we find that schools that are downward-shocked – receiving a lower accountability grade post-change than would have happened before – are less likely to retain their teachers than are schools that received no accountability shock. While downward-shocked schools in general lose more teachers, the effects are strongest for those shocked downward to a grade of "F". Teachers in schools who unexpectedly receive a grade of "F" are over 40 percent more likely to leave their school and are nearly 70 percent more likely to move to another school in the same district than are teachers in schools that did not receive an accountability shock. We also find some evidence the upward-shock schools are better able to retain their teachers than no-shock schools, though the measured effects are smaller and less precise than for downward-shock schools.

The mobility caused by school accountability also has significant effects on the distribution of teacher quality within and across schools. We find that downward accountability shocks lead to an increase in the quality of teachers who leave. However, increased accountability pressure also leads to improvements in the quality of teachers who remain in downwardly shocked schools. In contrast, schools that are not hit with an accountability shock do not experience any significant change to the quality of teachers that leave or stay. While the quality differential between stayers and leavers does not change as a result of increased accountability pressure, we do find that downward shocked schools experience changes to the distribution of the quality of their teachers. The upper tail of the quality distribution of teachers staying in negatively-shocked schools is truncated as a result of increased accountability

pressure. Further, the quality distribution of teachers who move shifts to the right while the distribution of those who stay remains relatively constant. It may be the case that the distribution of teachers staying or leaving schools shocked to "F" grades may be different, but we are hesitant to make those comparisons given the relatively small number of "F" school teachers for whom we have data to conduct the teacher quality analysis.

The results have strong implications for public policy. Struggling schools that come under increased accountability pressure face many challenges in terms of changing instructional policies and practices to facilitate student improvement. We have discovered that (in the case of those facing the highest accountability pressure) they also face the challenge of having to replace more teachers, and particularly, their higher-quality teachers (measured in terms of contribution to value-added). On the positive side, schools facing increased accountability pressure also see a rise in the average quality of the teachers who stay. It may be that those who stay respond work harder as a result of increased accountability pressure or they may appear to be more productive as a result of increases in complementary inputs like reading coaches. The findings presented in this paper suggest that if these schools were able to retain more of their high-quality teachers (perhaps through increased incentives to remain in the school), the accountability gains could be greater still. This last argument, of course, is still speculative.

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Table 1: The Distribution of School Grades, by Year

| School Grade | School Year | | | | | |
|--------------|-------------|-------------|-------------|-------------|-------------|-------------|
| | Summer 1999 | Summer 2000 | Summer 2001 | Summer 2002 | Summer 2003 | Summer 2004 |
| A | 183 | 552 | 570 | 887 | 1235 | 1203 |
| B | 299 | 255 | 399 | 549 | 565 | 515 |
| C | 1180 | 1115 | 1074 | 723 | 533 | 568 |
| D | 565 | 363 | 287 | 180 | 135 | 170 |
| F | 70 | 4 | 0 | 60 | 31 | 34 |
| N | 0 | 0 | 76 | 102 | 2 | 0 |
| Total | 2297 | 2289 | 2330 | 2501 | 2501 | 2490 |

Source: Authors' calculations from state data.

**Table 2: Transition Matrix in Predicted Grades Based on 2002 Grade Change
(row fractions)**

| | | Grade in 2002 based on new (summer 2002) grading system | | | | |
|---|---|---|---------------|---------------|---------------|--------------|
| | | A | B | C | D | F |
| Simulated grade in 2002 based on old (summer 2001) grading system | A | 0.90 [284] | 0.11 [33] | 0.00 [0] | 0.00 [0] | 0.00 [0] |
| | B | 0.59 [390] | 0.32 [209] | 0.09 [61] | 0.00 [1] | 0.00 [0] |
| | C | 0.18 [206] | 0.26 [301] | 0.49 [567] | 0.06 [69] | 0.00 [3] |
| | D | 0.02 [4] | 0.01 [3] | 0.37 [92] | 0.41 [103] | 0.20 [49] |
| | F | 0.00 [0] | 0.00 [0] | 0.00 [0] | 0.29 [2] | 0.71 [5] |

Notes: All row fractions are student-weighted. The number of schools is in brackets. Simulated grade changes are generated by applying both the old grading system and the new grading system to 2002 student test scores, using the approach introduced by Rouse et al. (2007). They are therefore generated based on precisely the same student tests; the only differences in the calculations are the formulas used to convert these same tests into school grades.

Table 3: Difference-in-Differences Estimates of the Impact of Accountability Shock on Fraction of Teachers Leaving School

| | Before Accountability Shock | After Accountability Shock | Time Difference for school |
|---|-----------------------------------|----------------------------------|-------------------------------|
| A. Comparison between Upward Shocked Schools and No Shock Schools | | | |
| Upward accountability shocked schools | 0.159 (0.001) [141924] | 0.166 (0.002) [49692] | 0.007 (0.002)*** |
| No accountability shocked schools | 0.166 (0.001) [167840] | 0.176 (0.002) [57865] | 0.010 (0.002)*** |
| Difference at a point in time | -0.007 (0.001)*** | -0.011 (0.002)*** | |
| Difference-in- differences | -0.004 (0.003) | | |
| B. Comparison between Downward Shocked Schools and No Shock Schools | | | |
| Downward accountability shocked schools | 0.182 (0.002) [28815] | 0.210 (0.004) [10215] | 0.028 (0.004)*** |
| No accountability shocked schools | 0.166 (0.001) [167840] | 0.176 (0.002) [57865] | 0.010 (0.002)*** |
| Difference at a point in time | 0.016 (0.002)*** | 0.033 (0.004)*** | |
| Difference-in- differences | 0.018 (0.005)** | | |

Note: Authors' calculations from state data. The fraction of teachers who left school is listed for each group. Standard errors clustered at the school-by-year level are given in the parentheses; sample sizes are given in square brackets; p<0.001***, p<0.01**, p<0.05*.

Table 4: Logit/Conditional Logit Estimates of the Effect of Accountability Shocks in Summer 2002 on the Odds of Teacher Job Change Between the 2002-03 and 2003-04 School Years

| | Logit | | | | | Conditional Logit | |
|--|------------------------------|------------------------------|------------------------------|------------------------------|------------------------------|------------------------------|------------------------------|
| | Model 1 | Model 2 | Model 3 | Model 4 | Model 5 | Model 5 | Model 6 |
| Effect of downward accountability shock (downward shocked schools after policy shock) | 1.111 (0.072) [p=0.10] | 1.110 (0.072) [p=0.11] | 1.116 (0.068) [p=0.07] | 1.124 (0.067) [p=0.05] | 1.121 (0.069) [p=0.06] | 1.103 (0.073) [p=0.14] | 0.981 (0.051) [p=0.71] |
| Effect of upward accountability shock (upward shocked schools after policy shock) | 0.977 (0.031) [p=0.48] | 0.975 (0.031) [p=0.42] | 0.971 (0.031) [p=0.35] | 0.972 (0.031) [p=0.36] | 0.965 (0.032) [p=0.29] | 0.962 (0.029) [p=0.20] | 0.935 (0.036) [p=0.08] |
| Effect of being downward shocked to grade F | | | | | | | 1.424 (0.220) [p=0.02] |
| Effect of being upward shocked to grade A | | | | | | | 1.051 (0.047) [p=0.27] |
| Year and post-shock indicators, shock-type fixed effect (logit) or school fixed effect (conditional logit) | ✓ | ✓ | ✓ | ✓ | ✓ | ✓ | ✓ |
| District-level demographics | | ✓ | ✓ | ✓ | ✓ | ✓ | ✓ |
| School-level demographics | | | ✓ | ✓ | ✓ | ✓ | ✓ |
| School-level FCAT math score | | | | ✓ | ✓ | ✓ | ✓ |
| School-level math score gains | | | | | ✓ | ✓ | ✓ |
| Number of teacher year observations | 456,342 | 456,342 | 438,245 | 438,071 | 384,637 | 384,598 | 384,598 |

Note: The parameter estimates presented above are expressed in terms of odds ratios relative to the unshocked set of teachers. Standard errors clustered at the school-by-year level are in parentheses beneath parameter estimates. The dependent variable for the logit models is the probability that a teacher will leave a school between year t and t+1. The logit models also include indicators for year, the post-shock period, an upward shocked school fixed effect, a downward shocked school fixed effect and school and district characteristics such as percent of minority students (Black, Hispanic), percent of students on free

or reduced lunch program, average math score on the Sunshine Standards FCAT, disciplinary incidents, and school-level math FCAT raw score gains. The dependent variable for conditional logit models are the probability that a teacher will leave a school within the school district between year t and $t+1$. In the conditional logit model school fixed effects replace the shock level fixed effects in logit model.

Table 5: Conditional Logit by Transition Destination Estimates of the Effect of Accountability Shocks in Summer 2002 on the Odds of Teacher Job Change Between the 2002-03 and 2003-04 School Years

| | Conditional Logit by Transition Destination | | |
|---|---|------------------------------|------------------------------|
| | Intra-district move | Inter-district move | Exit Florida public schools |
| Effect of downward accountability shock (downward shocked schools after policy shock) | 0.950 (0.081) [p=0.55] | 0.875 (0.106) [p=0.27] | 1.038 (0.059) [p=0.51] |
| Effect of upward accountability shock (upward shocked schools after policy shock) | 0.969 (0.065) [p=0.64] | 0.880 (0.069) [p=0.10] | 0.938 (0.038) [p=0.11] |
| Effect of being downward shocked to grade F | 1.665 (0.382) [p=0.03] | 1.058 (0.209) [p=0.78] | 1.167 (0.106) [p=0.09] |
| Effect of being upward shocked to grade A | 1.003 (0.081) [p=0.97] | 1.085 (0.096) [p=0.36] | 1.066 (0.049) [p=0.16] |
| Year and post-shock indicators, school fixed effects | ✓ | ✓ | ✓ |
| District-level demographics | ✓ | ✓ | ✓ |
| School-level demographics | ✓ | ✓ | ✓ |
| School-level FCAT math score | ✓ | ✓ | ✓ |
| School-level math score gains | ✓ | ✓ | ✓ |
| Number of teacher-year observations | 379,928 | 338,293 | 384,090 |

Note: The parameter estimates presented above are expressed in terms of odds ratios relative to the unshocked set of teachers. Standard errors clustered at the school-by-year level are in parentheses beneath parameter estimates. Conditional logit by transition destination models the probability of moving to a new school, moving to a new district, and exiting Florida public schools separately between year t and $t+1$. The number of observations may differ across these transition outcomes depending on the number of dropped teacher/year observation due to all positive or negative outcomes. The conditional logit models also include year indicators, post-shock indicators, school fixed effects, school and district characteristics such as percent of minority students (Black, Hispanic), percent of students on free or reduced lunch program, average math score on the Sunshine Standards FCAT, disciplinary incidents, and school-level math FCAT raw score gains. A full set of coefficient estimates is available upon request.

Table 6: Average Math Teacher Quality Differences for Stayers and Leavers across Zero-, Downward-, and Upward Shock Schools -- Before and After Policy Shock

| School/year | Before Accountability Shock | After Accountability Shock | Time Difference for School Type |
|---|-----------------------------|-----------------------------|---------------------------------|
| A. Comparison between Stayers and Leavers in Upward Shocked Schools | | | |
| Upward accountability shocked schools Leavers | 0.018 (0.009) [1004] | 0.005 (0.011) [735] | -0.013 (0.014) |
| Upward accountability shocked schools Stayers | 0.010 (0.002) [10149] | -0.001 (0.003) [5395] | -0.011 (0.004) |
| Difference at a point in time | 0.008 (0.008) | 0.006 (0.009) | |
| Difference-in-differences | -0.002 (0.013) | | |
| B. Comparison between Stayers and Leavers in Downward Shocked Schools | | | |
| Downward accountability shocked schools Leavers | -0.048 (0.021) [226] | 0.025 (0.022) [177] | 0.072 (0.031)** |
| Downward accountability shocked schools Stayers | -0.010 (0.007) [1730] | 0.036 (0.008) [907] | 0.046 (0.011)*** |
| Difference at a point in time | -0.036 (0.019)** | -0.012 (0.021) | |
| Difference-in-differences | 0.026 (0.030) | | |
| C. Comparison between Stayers and Leavers in No Accountability Shocked Schools | | | |
| No accountability shocked schools Leavers | -0.015 (0.008) [1185] | -0.019 (0.009) [890] | -0.004 (0.012) |
| No accountability shocked schools Stayers | 0.004 (0.002) [11215] | 0.002 (0.003) [5987] | -0.002 (0.004) |
| Difference at a point in time | -0.019 (0.008)*** | -0.022 (0.009)*** | |
| Difference-in-differences | -0.002 (0.012) | | |

Note: The teacher quality measure used here is obtained through a value-added specification with student's gain scores as the dependent variable. Other right hand side parameters are student covariates (gender, race/ethnicity, free/reduced-price lunch eligibility, gifted program participation, limited English proficiency, disability status, mobility), classroom peer characteristics (gender, race mobility, age, class size) and school-level covariates (new school indicator, experience of principal, indicator for new principal at school); p<0.001***, p<0.01**, p<0.05*

p<0.05*. #statistically significant at 10% level. **Table 7: Difference-in-Differences Estimates of the Impact of Accountability Shock on the Leavers' Average Math Quality**

| School/year | Before Accountability Shock | After Accountability Shock | Time Difference for school |
|---|-----------------------------------|----------------------------------|-------------------------------|
| D. Comparison between Upward Shocked Schools and No Shock Schools | | | |
| Upward accountability shocked schools | 0.018 (0.009) [1004] | 0.005 (0.011) [735] | -0.013 (0.014) |
| No accountability shocked schools | -0.015 (0.008) [1185] | -0.019 (0.009) [890] | -0.004 (0.012) |
| Difference at a point in time | 0.033 (0.012)*** | 0.025 (0.014)# | |
| Difference-in- differences | -0.009 (0.019) | | |
| E. Comparison between Downward Shocked Schools and No Shock Schools | | | |
| Downward accountability shocked schools | -0.048 (0.021) [226] | 0.025 (0.022) [177] | 0.072 (0.031)** |
| No accountability shocked schools | -0.015 (0.008) [1185] | -0.019 (0.009) [890] | -0.004 (0.012) |
| Difference at a point in time | -0.033 (0.020) | 0.044 (0.023)# | |
| Difference-in- differences | 0.077 (0.031)** | | |

Note: for definition of teacher quality measure, see note to Table 6; p<0.001***, p<0.01**, p<0.05*. #statistically significant at 10% level.

Table 8: Difference-in-Differences Estimates of the Impact of Accountability Shock on the Stayers' Average Math Quality

| School/year | Before Accountability Shock | After Accountability Shock | Time Difference for school |
|---|-----------------------------------|----------------------------------|-------------------------------|
| A. Comparison between Upward Shocked Schools and No Shock Schools | | | |
| Upward accountability shocked schools | 0.010 (0.002) [10149] | -0.001 (0.003) [5395] | -0.011 (0.004) |
| No accountability shocked schools | 0.004 (0.002) [11215] | 0.002 (0.003) [5987] | -0.002 (0.004) |
| Difference at a point in time | 0.006 (0.003)# | -0.003 (0.005) | |
| Difference-in- differences | -0.009 (0.006) | | |
| B. Comparison between Downward Shocked Schools and No Shock Schools | | | |
| Downward accountability shocked schools | -0.010 (0.007) [1730] | 0.036 (0.008) [907] | 0.046 (0.011)*** |
| No accountability shocked schools | 0.004 (0.002) [11215] | 0.002 (0.003) [5987] | -0.002 (0.004) |
| Difference at a point in time | -0.015 (0.007)* | 0.034 (0.009)*** | |
| Difference-in- differences | 0.048 (0.011)*** | | |

Note: for definition of teacher quality measure, see note to Table 6; p<0.001***, p<0.01**, p<0.05*. #statistically significant at 10% level.

Table 9: Difference-in-Differences Estimates of the Impact of Accountability Shock on the School Average Math Teacher Quality Difference between Stayers and Leavers

| School/year | Before Accountability Shock | After Accountability Shock | Time Difference for school |
|---|-----------------------------------|----------------------------------|-------------------------------|
| A. Comparison between Upward Shocked Schools and No Shock Schools | | | |
| Upward accountability shocked schools | -0.011 (0.011) [685] | -0.017 (0.014) [483] | -0.006 (0.017) |
| No accountability shocked schools | 0.015 (0.011) [801] | 0.021 (0.011) [575] | 0.006 (0.016) |
| Difference at a point in time | -0.026 (0.015)* | -0.039 (0.018)* | |
| Difference-in- differences | -0.012 (0.024) | | |
| B. Comparison between Downward Shocked Schools and No Shock Schools | | | |
| Downward accountability shocked schools | 0.042 (0.026) [145] | 0.041 (0.027) [107] | -0.001 (0.039) |
| No accountability shocked schools | 0.015 (0.011) [801] | 0.021 (0.011) [575] | 0.006 (0.016) |
| Difference at a point in time | 0.027 (0.027) | 0.019 (0.031) | |
| Difference-in- differences | -0.007 (0.041) | | |

Note: for definition of teacher quality measure, see note to Table 6; $p < 0.001^{***}$, $p < 0.01^{**}$, $p < 0.05^*$. #statistically significant at 10% level.

Figure 1: Kernel Density Estimates of Teacher Quality Distribution by School Type and Move Status, Before and After Accountability Shock

